

Q₁ - Let $A = \begin{bmatrix} 3 & -5 & 0 \\ 5 & -3 & 0 \\ 0 & 0 & 2 \end{bmatrix}$. Use the Gram-Schmidt

process to find an orthogonal basis for column of A.

- Solution:

$$A = \begin{bmatrix} 3 & -5 & 0 \\ 5 & -3 & 0 \\ 0 & 0 & 2 \end{bmatrix} \Rightarrow s_1 = \begin{bmatrix} 3 \\ 5 \\ 0 \end{bmatrix}, s_2 = \begin{bmatrix} -5 \\ -3 \\ 0 \end{bmatrix}, s_3 = \begin{bmatrix} 0 \\ 0 \\ 2 \end{bmatrix}$$

step 1) Let $\phi_1 = \frac{s_1}{\|s_1\|}$, $\|s_1\| = \sqrt{(3)^2 + (5)^2 + 0} = \sqrt{34}$

$$\Rightarrow \phi_1 = \frac{s_1}{\sqrt{34}} = \begin{bmatrix} \frac{3}{\sqrt{34}} \\ \frac{5}{\sqrt{34}} \\ 0 \end{bmatrix}$$

step 2)

$$(s_2)_{\perp s_1} = s_2 - (s_2^H \phi_1) \cdot \phi_1$$

$$s_2^H \phi_1 = \begin{bmatrix} -5 & -3 & 0 \end{bmatrix} \begin{bmatrix} \frac{3}{\sqrt{34}} \\ \frac{5}{\sqrt{34}} \\ 0 \end{bmatrix} = \frac{-15}{\sqrt{34}} - \frac{15}{\sqrt{34}} = \frac{-30}{\sqrt{34}}$$

$$\Rightarrow (s_2)_{\perp s_1} = \begin{bmatrix} -5 \\ -3 \\ 0 \end{bmatrix} - \left(\frac{-30}{\sqrt{34}} \right) \cdot \begin{bmatrix} \frac{3}{\sqrt{34}} \\ \frac{5}{\sqrt{34}} \\ 0 \end{bmatrix}$$

$$\Rightarrow (s_2)_{\perp s_1} = \begin{bmatrix} -5 \\ -3 \\ 0 \end{bmatrix} + \begin{bmatrix} \frac{45}{17} \\ \frac{75}{17} \\ 0 \end{bmatrix} = \begin{bmatrix} \frac{-85+45}{17} \\ \frac{-51+75}{17} \\ 0 \end{bmatrix} = \begin{bmatrix} \frac{-40}{17} \\ \frac{24}{17} \\ 0 \end{bmatrix}$$

$$\Rightarrow \phi_2 = \frac{(s_2)_{\perp s_1}}{\|(s_2)_{\perp s_1}\|}$$

we need to calculate $\|(s_2)_{\perp s_1}\|$:

$$\begin{aligned} \|(s_2)_{\perp s_1}\| &= \sqrt{\left(\frac{-40}{17}\right)^2 + \left(\frac{24}{17}\right)^2 + 0} = \frac{\sqrt{1600+576}}{17} = \frac{\sqrt{2176}}{17} \\ &= \frac{46.6}{17} \end{aligned}$$

$$\Rightarrow \phi_2 = \frac{(s_2)_{\perp s_1}}{\|(s_2)_{\perp s_1}\|} = \begin{bmatrix} \left(\frac{-40}{17}\right) / \left(\frac{46.6}{17}\right) \\ \left(\frac{24}{17}\right) / \left(\frac{46.6}{17}\right) \\ 0 \end{bmatrix} = \begin{bmatrix} \frac{-40}{46.6} \\ \frac{24}{46.6} \\ 0 \end{bmatrix}$$

step 3)

$$(s_3)_{\perp s_2} = s_3 - (s_3^H \phi_1) \cdot \phi_1 - (s_3^H \phi_2) \cdot \phi_2$$

$$s_3^H \phi_1 = [0 \ 0 \ 2] \begin{bmatrix} \frac{3}{\sqrt{34}} \\ \frac{5}{\sqrt{34}} \\ 0 \end{bmatrix} = 0, \quad s_3^H \phi_2 = [0 \ 0 \ 2] \begin{bmatrix} \frac{-40}{46.6} \\ \frac{24}{46.6} \\ 0 \end{bmatrix} = 0$$

$$\Rightarrow (s_3)_{\perp s_2} = s_3 \Rightarrow \phi_3 = \frac{(s_3)_{\perp s_2}}{\|(s_3)_{\perp s_2}\|} = \frac{s_3}{\|s_3\|} = \begin{bmatrix} 0 \\ 0 \\ 1 \end{bmatrix}$$

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Q₂ - Consider the set of three finite energy signals

$$s_1(t) = 1, \quad 0 \leq t \leq 1$$

$$s_2(t) = \cos(2\pi t), \quad 0 \leq t \leq 1$$

$$s_3(t) = \cos^2(\pi t), \quad 0 \leq t \leq 1$$

Find an orthonormal basis for the signal space spanned by these three signals?

- solution :

Firstly, we can write:

$$s_3(t) = \cos^2(\pi t) = \frac{1}{2} (1 + \cos(2\pi t)) = \frac{1}{2} (1) + \frac{1}{2} \cdot \cos(2\pi t)$$

$$= \frac{1}{2} s_1(t) + \frac{1}{2} \cos(2\pi t) = \frac{1}{2} s_1(t) + \frac{1}{2} s_2(t)$$

Thus, $s_3(t)$ can be written as a linear combination of $s_1(t)$ and $s_2(t)$. As a result, the dimension of the signal subspace is two instead of three which means we only need to perform the Gram-Schmidt procedure for two of the signals.

$$\text{step 1)} \quad v_1(t) = s_1(t) \quad , \quad \phi_1(t) = \frac{v_1(t)}{\|v_1(t)\|}$$

$$\|v_1(t)\|^2 = \int_0^1 v_1(t) v_1^*(t) dt = \int_0^1 1 dt = 1 \Rightarrow \|v_1(t)\| = 1$$

$$\Rightarrow \phi_1(t) = \frac{v_1(t)}{\|v_1(t)\|} = s_1(t) = 1 \quad 0 \leq t \leq 1$$

step 2)

$$v_2(t) = s_2(t) - \langle s_2(t), \phi_1(t) \rangle \cdot \phi_1(t)$$

$$\begin{aligned} \langle s_2(t), \phi_1(t) \rangle &= \int_0^1 s_2(t) \cdot \phi_1^*(t) dt = \int_0^1 \cos(2\pi t) \cdot (1) dt \\ &= \int_0^1 \cos(2\pi t) dt = \frac{1}{2\pi} [\sin(2\pi t)]_0^1 = \frac{1}{2\pi} [\sin(2\pi) - \sin(0)] \\ &= 0 \end{aligned}$$

$$\Rightarrow v_2(t) = s_2(t) = \cos(2\pi t), \quad \phi_2(t) = \frac{v_2(t)}{\|v_2(t)\|}$$

$$\|v_2(t)\|^2 = \int_0^1 \cos^2(2\pi t) dt = \int_0^1 \frac{1}{2} (1 + \cos(4\pi t)) dt$$

$$= \frac{1}{2} + \frac{1}{2} \int_0^1 \cos(4\pi t) dt = \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{4\pi} [\sin(4\pi t)]_0^1$$

$$= \frac{1}{2} + \frac{1}{2} \cdot \frac{1}{4\pi} [\sin(4\pi) - \sin(0)] = \frac{1}{2}$$

$$\Rightarrow \|v_2(t)\| = \frac{1}{\sqrt{2}}$$

$$\phi_2(t) = \frac{\cos(2\pi t)}{(\frac{1}{\sqrt{2}})} = \sqrt{2} \cdot \cos(2\pi t)$$

Q3 - Show the following statement is true for x being a normal random variable:

$$\int_{-\infty}^{+\infty} f_x(x) dx = 1$$

- solution:

$$\int_{-\infty}^{+\infty} f_x(x) dx = \frac{1}{\sqrt{2\pi} \sigma} \int_{-\infty}^{+\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \quad (1)$$

$$\text{Let } y = \frac{(x-\mu)}{\sqrt{2} \sigma} \Rightarrow dx = \sqrt{2} \sigma dy \quad (2)$$

$$\Rightarrow (1) = \frac{1}{\sqrt{2\pi} \sigma} \int_{-\infty}^{+\infty} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx \xrightarrow{\text{substitute (2)}} \frac{1}{\sqrt{\pi}} \int_{-\infty}^{+\infty} e^{-y^2} dy$$

$$\text{Let } \int_{-\infty}^{+\infty} e^{-y^2} dy = I$$

$$\text{Then } I^2 = \left[\int_{-\infty}^{+\infty} e^{-x^2} dx \right] \left[\int_{-\infty}^{+\infty} e^{-y^2} dy \right]$$

$$= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} e^{-(x^2+y^2)} dx dy$$

Letting $x = r \cdot \cos(\theta)$, $y = r \cdot \sin(\theta)$ (polar coordinates)
we have:

$$I^2 = \int_0^{2\pi} \int_0^{\infty} e^{-r^2} r \cdot dr d\theta = 2\pi \int_0^{\infty} e^{-r^2} \cdot r \cdot dr =$$

$$2\pi \cdot \left[-\frac{1}{2} e^{-r^2} \right]_0^{\infty} = -\pi \left[e^{-r^2} \right]_0^{\infty} = -\pi [0 - 1] = \pi$$

$$\Rightarrow I = \sqrt{\pi} \Rightarrow \int_{-\infty}^{+\infty} f_x(x) dx = \frac{1}{\sqrt{\pi}} I = \frac{1}{\sqrt{\pi}} \cdot \sqrt{\pi} = 1 \quad \checkmark$$

Q4 - show that the conditional pdf of x
given the event $B = (a < x \leq b)$ is as follows:

$$f_x(x | a < x \leq b) = \begin{cases} 0 & x \leq a \\ \frac{f_x(x)}{\int_a^b f_x(x') dx'} & a < x \leq b \\ 0 & x > b \end{cases}$$

- solution:

$$F_x(x | a < x \leq b) = P(x \leq x | a < x \leq b)$$

$$= \frac{P\{(x \leq x) \cap (a < x \leq b)\}}{P(a < x \leq b)}$$

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$$\text{Now } (x \leq n) \cap (a < x \leq b) = \begin{cases} \emptyset & x \leq a \\ a < x \leq x & a < x \leq b \\ a < x \leq b & x > b \end{cases}$$

Thus,

$$F_x(x | a < x \leq b) = \frac{P(\emptyset)}{P(a < x \leq b)} = 0 \quad \text{for } x \leq a$$

$$F_x(x | a < x \leq b) = \frac{P(a < x \leq x)}{P(a < x \leq b)} = \frac{F_x(x) - F_x(a)}{F_x(b) - F_x(a)}, \quad a < x \leq b$$

$$F_x(x | a < x \leq b) = \frac{P(a < x \leq b)}{P(a < x \leq b)} = 1, \quad \text{for } x > b$$

$$\text{Now, we use } f_x(x) = \frac{d F_x(x)}{dx}$$

$$\Rightarrow f_x(x | a < x \leq b) = \begin{cases} 0 & x \leq a \\ \frac{f_x(x)}{F_x(b) - F_x(a)} = \frac{f_x(x)}{\int_a^b f_x(x') dx'} & a < x \leq b \\ 0 & x > b \end{cases}$$

Q5 - If $\underline{x} \sim \mathcal{N}(\underline{\mu}, \underline{\Sigma})$ and $\underline{A}$ is a $q \times p$ matrix with $\text{rank}(\underline{A}) = q \leq p$, show that

$$\underline{\mu}_z = \underline{A} \underline{\mu} + \underline{b} \quad \text{and} \quad \underline{\Sigma}_z = \underline{A} \underline{\Sigma} \underline{A}^T \quad \text{in the case that} \quad \underline{z} = \underline{A} \underline{x} + \underline{b}.$$

- solution:

Simply using the definition of mean and covariance matrix, we can write:

$$\begin{aligned} \underline{\mu}_z &= E\{\underline{z}\} = E\{\underline{A} \underline{x} + \underline{b}\} = E\{\underline{A} \underline{x}\} + E\{\underline{b}\} \\ &= \underline{A} E\{\underline{x}\} + \underline{b} = \underline{A} \underline{\mu} + \underline{b} \quad \checkmark \end{aligned}$$

$$\begin{aligned} \underline{\Sigma}_z &= E\{(\underline{z} - \underline{\mu}_z)(\underline{z} - \underline{\mu}_z)^T\} \\ &= E\{(\underline{A} \underline{x} + \underline{b} - (\underline{A} \underline{\mu} + \underline{b}))(\underline{A} \underline{x} + \underline{b} - (\underline{A} \underline{\mu} + \underline{b}))^T\} \\ &= E\{(\underline{A}(\underline{x} - \underline{\mu}))(\underline{A}(\underline{x} - \underline{\mu}))^T\} = E\{\underline{A}(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^T \underline{A}^T\} \\ &= \underline{A} \cdot \underbrace{E\{(\underline{x} - \underline{\mu})(\underline{x} - \underline{\mu})^T\}}_{\underline{\Sigma}} \underline{A}^T = \underline{A} \cdot \underline{\Sigma} \underline{A}^T \quad \checkmark \end{aligned}$$