



Aalto University
School of Electrical
Engineering

ELEC-E7450
Performance Analysis

Single server queue M/G/1

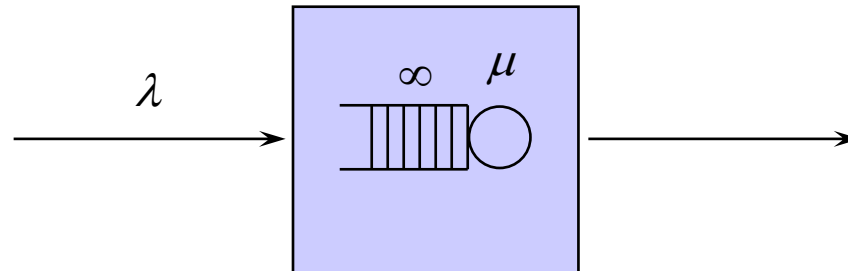
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Contents

- M/G/1 with a work-conserving service discipline
- M/G/1-FIFO
- M/G/1-PS
- Performance comparison between FIFO and PS

M/G/1

- Customers arrive according to a Poisson process at rate λ
 - IID inter-arrival times
 - **exponential inter-arrival time distribution** with mean $1/\lambda$
- Customers are served by **1** server
 - IID service times S_i
 - **general service time distribution** with mean $E[S] = 1/\mu$
- There are ∞ customer places in the system



Service discipline

- **Definition:**

Service discipline π determines the way the customers are served

- It specifies whether the customers are served one-by-one or simultaneously
- If the customers are served one-by-one, it specifies the order in which they are taken to service
- If the customers are served simultaneously, it specifies how the service capacity is shared among them

- Service discipline is also called as **queueing discipline**, or **scheduling discipline**

- **Definition:**

A service discipline is **work-conserving** if customers are served whenever the system is non-empty

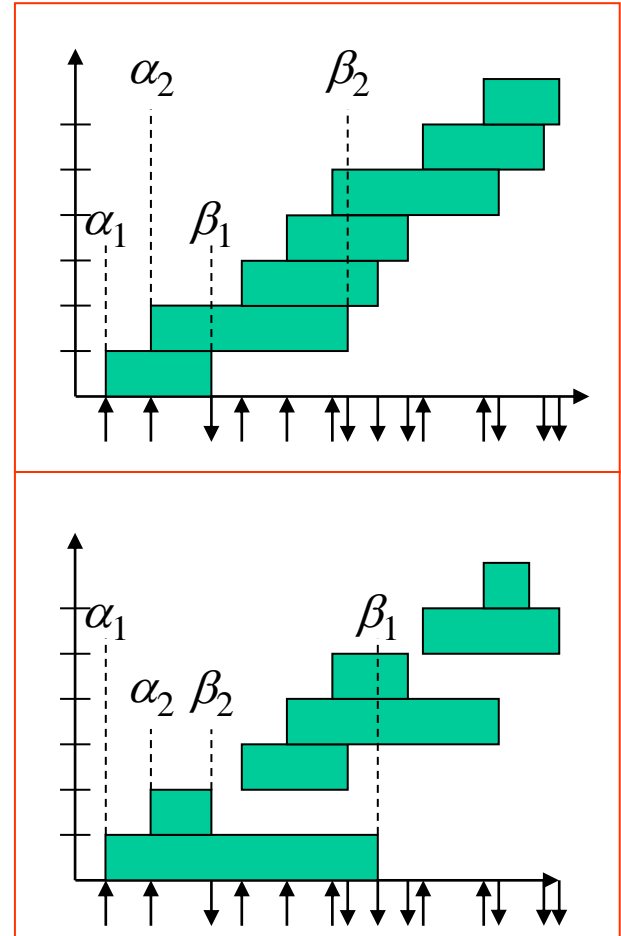
Work-conserving service disciplines

- **First In First Out (FIFO)**
 - customers are served one-by-one until completion
 - service in the arrival order (“ordinary queue”)
 - the customer that arrived first is served with rate μ
 - also known as **First Come First Served (FCFS)**
- **Processor Sharing (PS)**
 - customers are served simultaneously
 - the service capacity is shared evenly among all customers (“fair queue”)
 - when i customers are in the system, each of them is served with rate μ/i
 - ideal version of the **Round Robin (RR)** service discipline

Delay

- α_i = arrival time of customer i
- β_i^π = departure time of customer i
- T_i^π = delay (sojourn time) of customer i

$$T_i^\pi := \beta_i^\pi - \alpha_i$$



Queue length

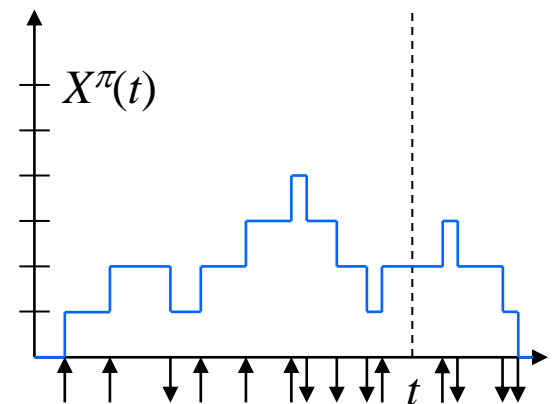
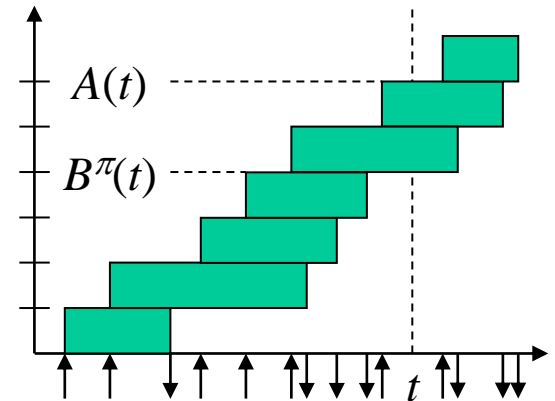
- $A(t)$ = number of arrivals until time t
= arrival process
- $B^\pi(t)$ = number of departures until time t
= departure process

$$A(t) := \max \{i : \alpha_i \leq t\}$$

$$B^\pi(t) := |\{i : \beta_i^\pi \leq t\}|$$

- $X^\pi(t)$ = number of customers at time t
= queue length process

$$X^\pi(t) := A(t) - B^\pi(t)$$



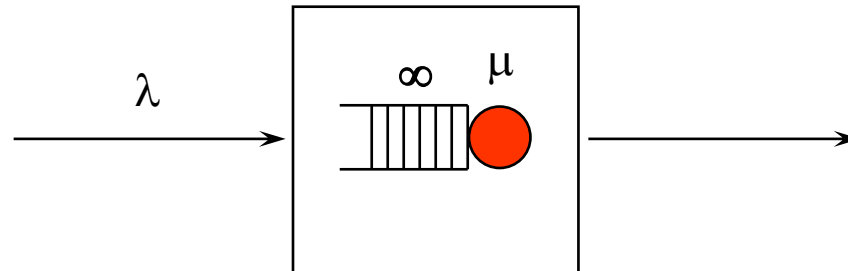
Traffic load in M/G/1

- **Definition:**
Traffic load ρ is defined by

$$\rho := \frac{\lambda}{\mu}$$

- Applying Little's formula to the subsystem consisting of just the server:

$$P\{X > 0\} = E[1_{\{X > 0\}}] \stackrel{\text{Little}}{=} \lambda E[S] = \frac{\lambda}{\mu} = \rho, \quad P\{X = 0\} = 1 - \rho$$



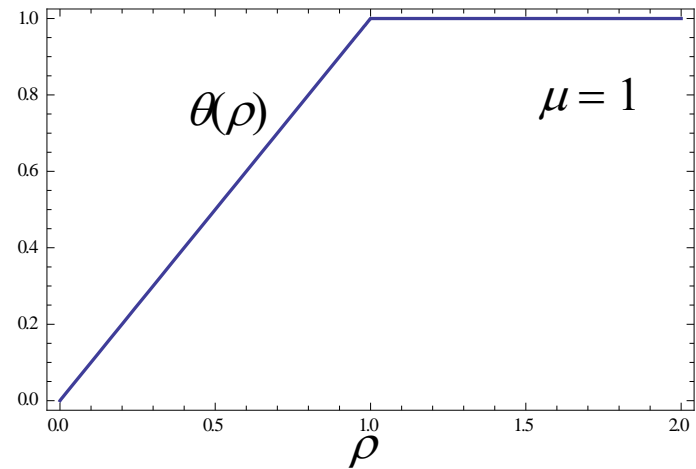
Throughput

- **Definition:**
Throughput θ of service discipline π refers to the long-run average departure rate given by

$$\theta := \lim_{t \rightarrow \infty} \frac{1}{t} E[B^\pi(t)]$$

- **Proposition:**
For any work-conserving service discipline π , the throughput is

$$\theta = \min\{\lambda, \mu\}$$



- **Corollary:**
 - If $\rho \leq 1$, then $\theta = \lambda$
 - If $\rho \geq 1$, then $\theta = \mu$

Stability

- **Definition:**

Service discipline π is **unstable** if

$$\lim_{t \rightarrow \infty} P\{X^\pi(t) \geq n\} = 1 \text{ for all } n$$

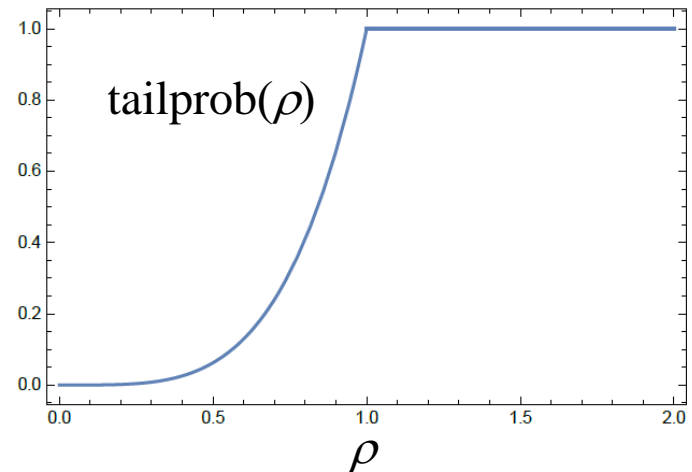
Otherwise it is **stable**.

- **Proposition:**

All service disciplines π are unstable if $\rho \geq 1$

- **Proposition:**

All work-conserving service disciplines π are stable if $\rho < 1$



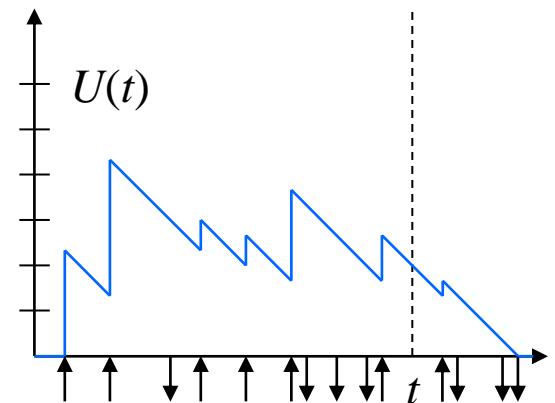
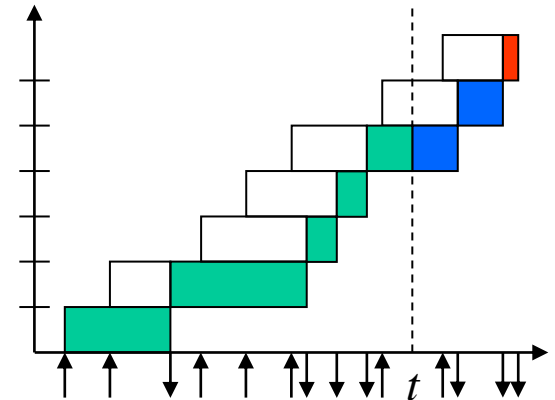
Unfinished work

- $U(t)$ = sum of remaining service times of all customers in system at time t
 = total **workload** at time t
 = **unfinished work process**

$$U(t) := \sum_{i=1}^{A(t)} S_i - \int_0^t 1_{\{U(s) > 0\}} ds$$

- **Proposition:**
 The unfinished work process $U(t)$ is the same for all work-conserving disciplines π . In addition,

$$\{X^\pi(t) = 0\} = \{U(t) = 0\}$$

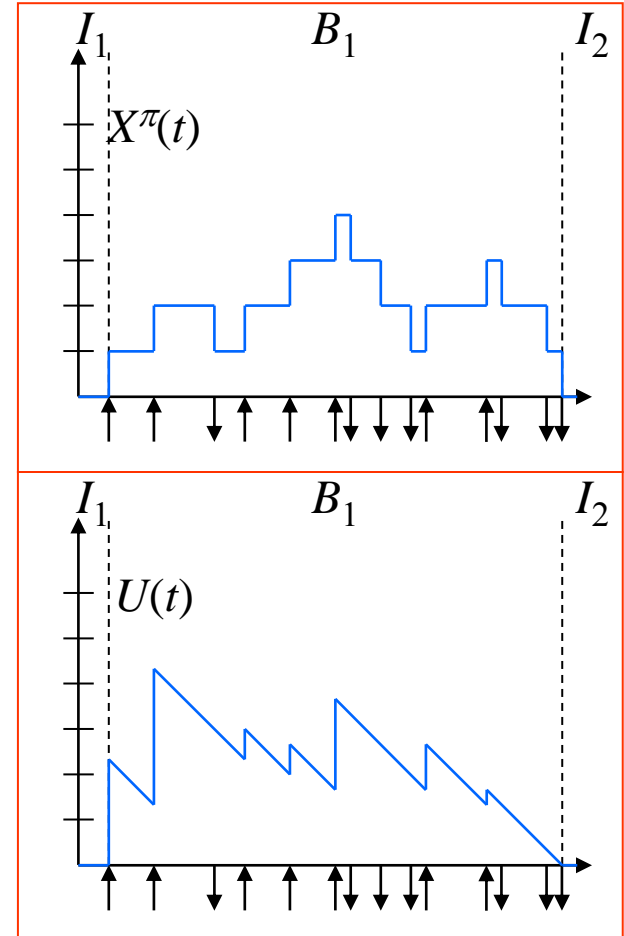


Busy and idle periods (1)

- Definition:**
 The server is **busy** whenever the system is non-empty, and **idle** otherwise.
 A **busy [idle] period** is an unbroken interval during which the server is busy [idle].

- I_n = length of n th idle period
- B_n = length of n th busy period
- C_n = length of n th busy cycle
- N_n = number of customers served in the n th busy period

$$C_n = I_n + B_n$$



Busy and idle periods (2)

- Assume that $U(0) = 0$.

$$\gamma_1 := 0$$

$$I_1 := \inf\{t > 0 \mid U(t) > 0\} = \alpha_1$$

$$B_1 := \inf\{t > I_1 \mid U(t) = 0\} - I_1$$

$$C_1 := I_1 + B_1$$

$$N_1 := A(C_1)$$

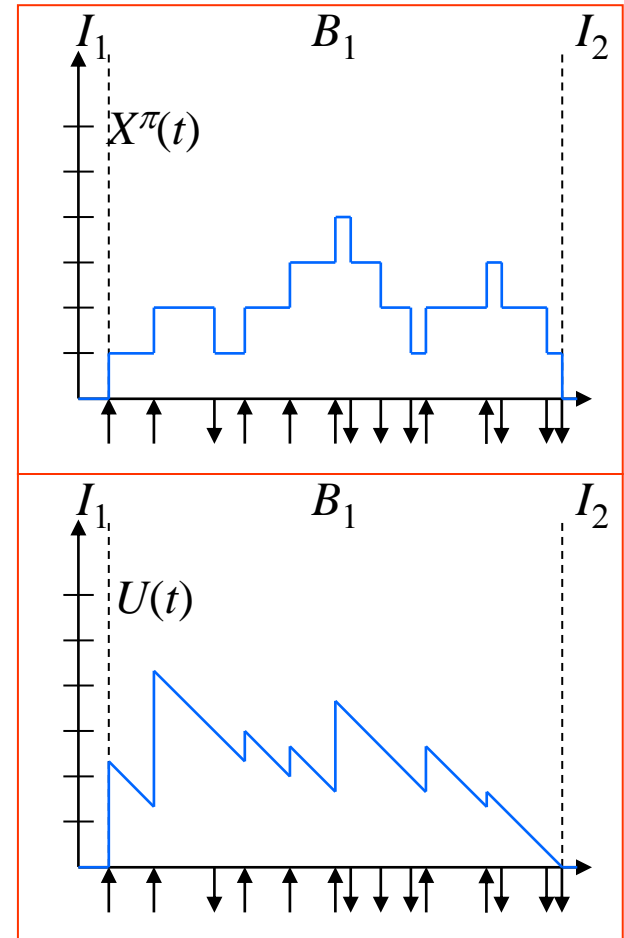
$$\gamma_n := \gamma_{n-1} + C_{n-1}$$

$$I_n := \inf\{t > \gamma_n \mid U(t) > 0\} - \gamma_n$$

$$B_n := \inf\{t > \gamma_n + I_n \mid U(t) = 0\} - \gamma_n - I_n$$

$$C_n := I_n + B_n$$

$$N_n := A(\gamma_n + C_n) - A(\gamma_n)$$



Busy and idle periods (3)

- **Proposition:**

Idle periods I_n , busy periods B_n , busy cycles C_n , and the number of customers N_n served in a busy period are the same for all work-conserving service disciplines π .

- **Proposition:**

The busy cycles C_n constitute a **renewal sequence** (γ_n) . In addition, $X^\pi(t)$ and $U(t)$ are **regenerative processes** with respect to the renewal sequence (γ_n) for all work-conserving service disciplines π .

- **Proposition:** Assume that $\rho < 1$.

(i) Idle periods I_n are IID with mean

$$E[I] = 1/\lambda$$

(ii) Busy periods B_n are IID with mean

$$E[B] = \frac{E[S]}{1-\rho}$$

(iii) Busy cycles C_n are IID with mean

$$E[C] = \frac{1/\lambda}{1-\rho}$$

(iv) Number of customers N_n served in a busy period are IID with mean

$$E[N] = \frac{1}{1-\rho}$$

Busy and idle periods (4)

- By previous propositions, the steady-state variables X^π and U ,

$$P\{X^\pi \leq x\} := \lim_{t \rightarrow \infty} P\{X^\pi(t) \leq x\}$$

$$= \frac{E[\int_0^C 1\{X^\pi(t) \leq x\} dt]}{E[C]}$$

$$P\{U \leq x\} := \lim_{t \rightarrow \infty} P\{U(t) \leq x\}$$

$$= \frac{E[\int_0^C 1\{U(t) \leq x\} dt]}{E[C]}$$

are well-defined whenever the system is stable, $\rho < 1$.

- Proposition:**

Assume that $\rho < 1$. For all work-conserving service disciplines π , we have

$$P\{X^\pi = 0\} = P\{U = 0\} = 1 - \rho$$

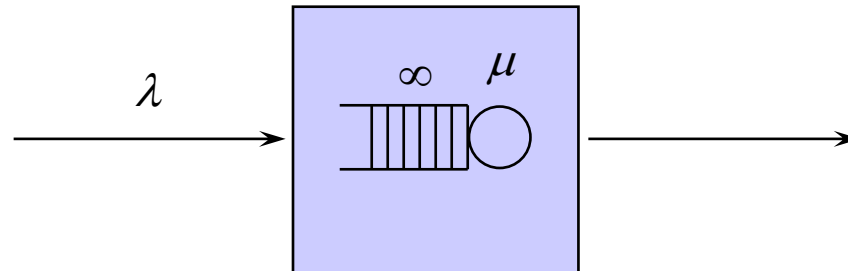
$$P\{X^\pi > 0\} = P\{U > 0\} = \rho$$

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M/G/1-FIFO

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FIFO service discipline

- **First In First Out (FIFO)**
 - customers are served one-by-one until completion
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 - also known as **First Come First Served (FCFS)**

Waiting time

- In a FIFO system, the delay T_i of customer i consists of its waiting time W_i and service time S_i

$$T_i = W_i + S_i$$

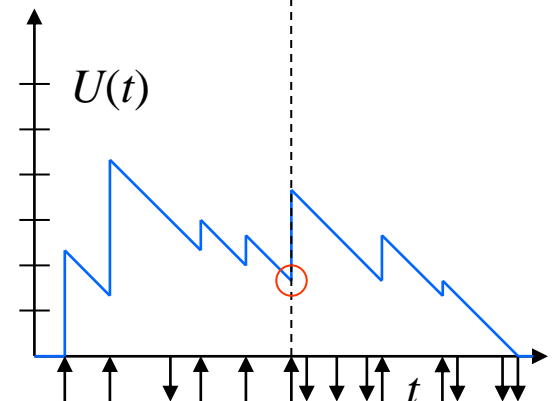
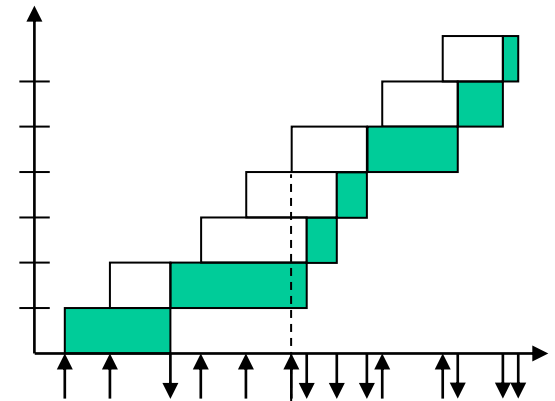
- Let Y_i^w denote the number of **waiting customers** that customer i sees upon its arrival,

$$Y_i^w = \max\{X(\alpha_i^-) - 1, 0\}$$

- Now

$$W_i = U(\alpha_i^-) = \sum_{j=1}^{Y_i^w} S_{i-j} + R(\alpha_i^-)$$

where $R(t)$ denotes the **remaining service time** of the customer in service at time t (if any).



Remaining service time

- **Proposition:**
Process $R(t)$ is **regenerative** with respect to the renewal sequence (γ_n) .

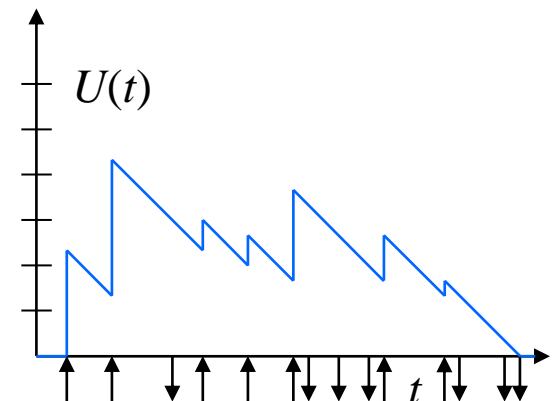
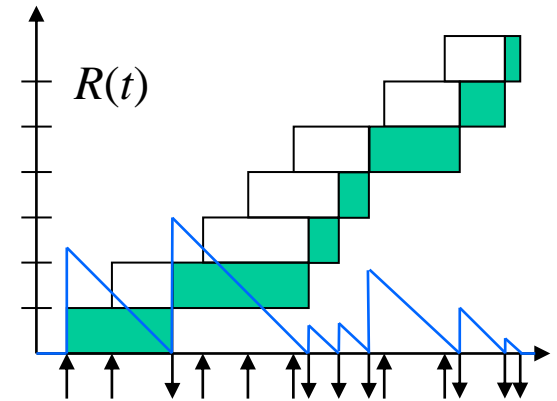
- Thus, the steady-state variable R ,

$$P\{R \leq x\} := \lim_{t \rightarrow \infty} P\{R(t) \leq x\} = \frac{E[\int_0^C 1\{R(t) \leq x\} dt]}{E[C]}$$

is well-defined whenever $\rho < 1$.

- **Proposition:**
Assume $\rho < 1$. Then

$$E[R] = \frac{\lambda}{2} E[S^2]$$



Pollaczek-Khinchin mean value formulas

- Theorem:**

Assume $\rho < 1$. For the M/G/1-FIFO queue, we have

$$E[W] = E[U] = \frac{\lambda E[S^2]}{2(1-\rho)}$$

$$E[T] = E[S] + \frac{\lambda E[S^2]}{2(1-\rho)}$$

$$E[X] = \rho + \frac{\lambda^2 E[S^2]}{2(1-\rho)}$$

- These are called the **Pollaczek-Khinchin mean value formulas** for M/G/1-FIFO

- Note that, if the mean service time $E[S]$ is kept fixed, then $E[W]$, $E[T]$, and $E[X]$ are increasing functions of the coefficient of variation $C[S]$ of the service time distribution,

$$C[S] := \sqrt{\frac{D^2[S]}{E[S]^2}} = \sqrt{\frac{E[S^2]}{E[S]^2} - 1}$$

$$\Rightarrow E[S^2] = E[S]^2 (1 + C^2[S])$$

- Examples:**

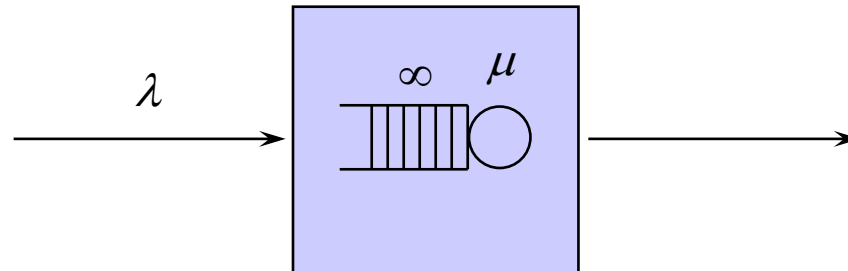
- Erlang distribution: $C[S] < 1$
- Exponential distribution: $C[S] = 1$
- Hyperexponential distrib.: $C[S] > 1$
- Pareto distribution: $C[S] > 1$

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PS service discipline

- **Processor Sharing (PS)**
 - customers are served simultaneously
 - the service capacity is shared evenly among all customers (“fair queue”)
 - when i customers are in the system, each of them is served with rate μ/i
 - ideal version of the **Round Robin (RR)** service discipline

Exponential distribution

$$X \sim \text{Exp}(\mu), \quad \mu > 0$$

$$S = (0, \infty)$$

$$F_X(x) := P\{X \leq x\} = 1 - e^{-\mu x}$$

$$f_X(x) := \frac{d}{dx} F_X(x) = \mu e^{-\mu x}$$

$$E[X] = \frac{1}{\mu}$$

$$E[X^2] = \frac{2}{\mu^2}$$

$$D^2[X] = \frac{1}{\mu^2}$$

$$D[X] = \frac{1}{\mu}$$

$$C[X] := \frac{D[X]}{E[X]} = 1$$



M/M/1-PS

- Let us first consider the **M/M/1-PS** queue
- So we assume that the service times obey the **Exp(μ)** distribution
- In this case, the queue length process **$X(t)$** is an **irreducible (Markov) birth-death process** with state space

$$S = \{0,1,2,\dots\}$$

and transition rates

$$q(n, n+1) = \lambda$$

$$q(n+1, n) = (n+1) \frac{\mu}{n+1} = \mu$$

- **Proposition:**

Assume $\rho < 1$. For the M/M/1-PS queue, the steady-state queue length distribution is

$$P\{X = n\} = (1 - \rho)\rho^n, \quad n \in \{0,1,2,\dots\}$$

Erlang distribution

- IID **exponential phases** in a series

$$X \sim \text{Erl}(K, K\mu), \quad \mu > 0$$

$$X = X_1 + \dots + X_K, \quad X_k \sim \text{Exp}(K\mu)$$

$$S = (0, \infty)$$

$$F_X(x) = 1 - \sum_{k=1}^K \frac{(K\mu x)^{K-k}}{(K-k)!} e^{-K\mu x}$$

$$f_X(x) = K\mu \frac{(K\mu x)^{K-1}}{(K-1)!} e^{-K\mu x}$$

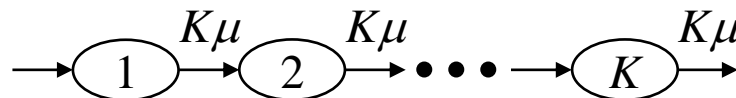
$$E[X] = \frac{1}{\mu}$$

$$E[X^2] = \frac{K+1}{K\mu^2}$$

$$D^2[X] = \frac{1}{K\mu^2}$$

$$D[X] = \frac{1}{\sqrt{K}\mu}$$

$$C[X] := \frac{D[X]}{E[X]} = \frac{1}{\sqrt{K}}$$



M/E_K/1-PS (1)

- Next we consider the **M/E_K/1-PS** queue
- Here we assume that the service times obey the **Erl(K, Kμ)** distribution with **K ≥ 2 exponential phases**
- In this case, the queue length process **X(t)** is no longer a Markov process, but we have to **supplement** the state description.
- To get a Markovian description of the system, we have to additionally keep track of the current phases of the customers.

- Let

$$N(t) = (N_1(t), \dots, N_K(t))$$

where **N_k(t)** refers to the **total number of customers in phase k** at time **t**

- Process **N(t)** is an **irreducible Markov process** with state space

$$S = \{n = (n_1, \dots, n_K) \mid n_k \in \{0, 1, 2, \dots\}\}$$

and transition rates

$$q(n, n + e_1) = \lambda$$

$$q(n + e_k, n + e_{k+1}) = \frac{(n_k + 1)K\mu}{n_1 + \dots + n_K + 1}, \quad k < K$$

$$q(n + e_K, n) = \frac{(n_K + 1)K\mu}{n_1 + \dots + n_K + 1}$$

M/E_K/1-PS (2)

- Proposition:**

Assume $\rho < 1$. The steady-state distribution of process $N(t)$ is

$$P\{N = n\} = (1 - \rho)(\rho / K)^{n_1 + \dots + n_K} \times \frac{(n_1 + \dots + n_K)!}{n_1! \dots n_K!}, \quad n \in \mathcal{S}$$

- Corollary:**

Assume $\rho < 1$. For the M/E_K/1-PS queue, the steady-state queue length distribution is

$$P\{X = n\} = (1 - \rho)\rho^n, \quad n \in \{0, 1, 2, \dots\}$$

- Note that

$$X(t) = N_1(t) + \dots + N_K(t)$$

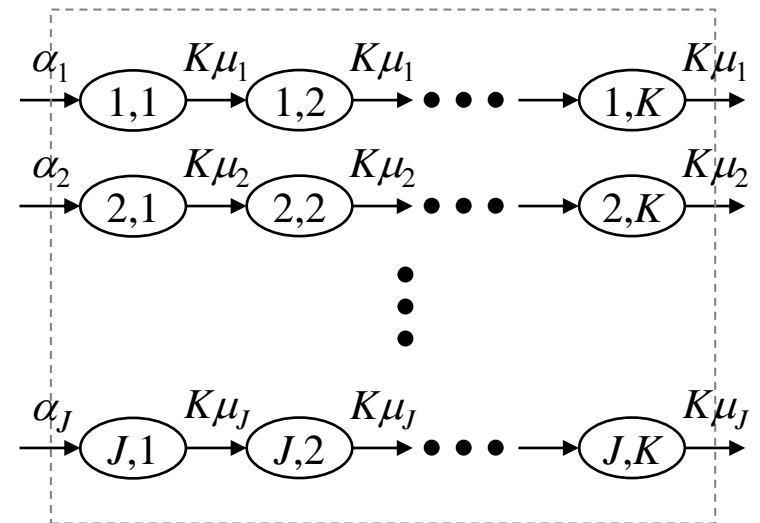
Phase-type distribution

- **Definition:**

A **phase-type (PH)** distribution refers to the distribution of the absorption time in an absorbing finite-state Markov process

- **Examples:**

- Exponential distribution
- Erlang distribution
- Hyperexponential distribution
- Exponential distributions in series and/or parallel



M/PH/1-PS (1)

- Next we consider the **M/PH/1-PS** queue
- Here we assume (for simplicity) that the service times obey the **phase-type** distribution depicted in the previous slide with $JK \geq 2$ **exponential phases**
- In this case, the queue length process $X(t)$ is neither a Markov process, but we have to **supplement** the state description as before.
- Again, to get a Markovian description of the system, we have to additionally keep track of the current phases of the customers.

- Let

$$N(t) = (N_{1,1}(t), \dots, N_{J,K}(t))$$

where $N_{j,k}(t)$ refers to the **total number of customers in phase (j,k)** at time t

- Process $N(t)$ is an **irreducible Markov process** with state space

$$S = \{n = (n_{1,1}, \dots, n_{J,K}) \mid n_{j,k} \in \{0, 1, 2, \dots\}\}$$

and transition rates determined from the underlying absorbing Markov process

- **Exercise:**
Determine the transition rates

M/PH/1-PS (2)

- Proposition:**

Assume $\rho < 1$. The steady-state distribution of process $N(t)$ is

$$P\{N = n\} = (1 - \rho) \times \prod_{j=1}^J (\rho_j / K)^{n_{j,1} + \dots + n_{j,K}} \times \frac{(n_{1,1} + \dots + n_{J,K})!}{n_{1,1}! \dots n_{J,K}!}, \quad n \in S$$

where

$$\rho := \rho_1 + \dots + \rho_J, \quad \rho_j := \frac{\lambda \alpha_j}{\mu_j}$$

- Note that

$$X(t) = N_{1,1}(t) + \dots + N_{J,K}(t)$$

- Corollary:**

Assume $\rho < 1$. For the M/PH/1-PS queue, the steady-state queue length distribution is

$$P\{X = n\} = (1 - \rho) \rho^n, \quad n \in \{0, 1, 2, \dots\}$$

Insensitive queue length distribution in M/G/1-PS

- The generalization of the previous result is based on the known fact that any service time distribution can be approximated (with an arbitrary precision) by a phase-type distribution.
- Since the queue length distribution remains the same for any service time distribution with the same mean $E[S]$, the steady-state queue length distribution of the PS service discipline is said to be **insensitive** to the service time distribution.
- Interestingly, the mean sojourn time $E[T]$ in the M/G/1-PS queue equals the mean busy period $E[B]$.

- **Theorem:**

Assume $\rho < 1$. For the M/G/1-PS queue, the steady-state queue length distribution is

$$P\{X = n\} = (1 - \rho)\rho^n, \quad n \in \{0, 1, 2, \dots\}$$

with

$$E[X] = \frac{\rho}{1 - \rho}$$

- **Corollary:**

Assume $\rho < 1$. For the M/G/1-PS queue, the mean steady-state sojourn time is

$$E[T] = \frac{E[S]}{1 - \rho}$$

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Performance comparison between FIFO and PS

- Proposition:**

Assume $\rho < 1$. For the M/G/1 queue, we have

$$E[X^{\text{FIFO}}] < E[X^{\text{PS}}] \Leftrightarrow E[T^{\text{FIFO}}] < E[T^{\text{PS}}] \Leftrightarrow C[S] < 1$$

$$E[X^{\text{FIFO}}] = E[X^{\text{PS}}] \Leftrightarrow E[T^{\text{FIFO}}] = E[T^{\text{PS}}] \Leftrightarrow C[S] = 1$$

$$E[X^{\text{FIFO}}] > E[X^{\text{PS}}] \Leftrightarrow E[T^{\text{FIFO}}] > E[T^{\text{PS}}] \Leftrightarrow C[S] > 1$$

where $C[S]$ refers to the coefficient of variation of the service time distribution

Summary

- **M/G/1 with a work-conserving service discipline**
 - traffic load, throughput, stability, unfinished work, busy period, busy cycle
- **M/G/1-FIFO**
 - FIFO, remaining service time, Pollaczek-Khinchin mean value formulas
- **M/G/1-PS**
 - PS, phase method, insensitivity
- **Performance comparison between FIFO and PS**
 - FIFO better than PS if the service time distribution less variable than exp