

# Random graphs and network statistics

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# Appendix A

## Probability

Here are some miscellaneous facts from probability theory that are used in the text.

### A.1 Inequalities

**Proposition A.1.1** (Markov's inequality). *For any random number  $X \geq 0$  and any  $a > 0$ ,*

$$\mathbb{P}(X \geq a) \leq a^{-1}\mathbb{E}X.$$

*Proof.* First, note that  $\mathbb{P}(X \geq a) = \mathbb{E}1(X \geq a)$  where  $1(A)$  in general denotes the *indicator* of the event  $A$ . Hence by the linearity of expectation,

$$a\mathbb{P}(X \geq a) = a\mathbb{E}1(X \geq a) = \mathbb{E}a1(X \geq a).$$

Next, the inequalities

$$a1(X \geq a) \leq X1(X \geq a) \leq X$$

which are valid for any realization of  $X$ , and the monotonicity of the expectation imply that

$$\mathbb{E}a1(X \geq a) \leq \mathbb{E}X.$$

Hence  $a\mathbb{P}(X \geq a) \leq \mathbb{E}X$ , and the claim follows.  $\square$

**Proposition A.1.2** (Chebyshev's inequality). *For any random number  $X$  with a finite mean  $\mu = \mathbb{E}X$  and any  $a > 0$ ,*

$$\mathbb{P}(|X - \mu| \geq a) \leq a^{-2} \text{Var}(X).$$

*Proof.* By applying Markov's inequality for  $Y = (X - \mu)^2$ , we find that

$$\begin{aligned}\mathbb{P}(|X - \mu| \geq a) &= \mathbb{P}((X - \mu)^2 \geq a^2) \\ &\leq (a^2)^{-1} \mathbb{E}(X - \mu)^2 = a^{-2} \text{Var}(X).\end{aligned}$$

□

The following inequality is due to the Finnish-born Wassily Hoeffding.

**Proposition A.1.3.** *Let  $S_n = \sum_{i=1}^n X_i$  where the summands are independent and bounded by  $a_i \leq X_i \leq b_i$ . Then for any  $t > 0$ ,*

$$\mathbb{P}(S_n \geq \mathbb{E}S_n + t) \leq e^{-\frac{2t^2}{\sum_i (b_i - a_i)^2}},$$

$$\mathbb{P}(S_n \leq \mathbb{E}S_n - t) \leq e^{-\frac{2t^2}{\sum_i (b_i - a_i)^2}},$$

and

$$\mathbb{P}(|S_n - \mathbb{E}S_n| \geq t) \leq 2e^{-\frac{2t^2}{\sum_i (b_i - a_i)^2}}.$$

*Proof.* A well-written proof of the first inequality, based on an extremality property related to convex stochastic orders, is available in the original research article Hoeffding [5]. The second inequality follows by applying the first inequality to  $\tilde{S}_n = -S_n$  and the third inequality follows from the first two by the union bound. □

## A.2 Weak convergence of probability measures

Let  $\mu, \mu_1, \mu_2, \dots$  be probability distributions on  $\mathbb{R}$ . We say that  $\mu_n \rightarrow \mu$  *weakly* if  $\int \phi(x) \mu_n(dx) \rightarrow \int \phi(x) \mu(dx)$  for every bounded continuous function  $\phi : \mathbb{R} \rightarrow \mathbb{R}$ . We say that  $\mu_n \rightarrow \mu$  *weakly and with  $k$ -th moments*, if in addition  $\mu_n$  and  $\mu$  have finite  $k$ -th moments and  $\int |x|^k \mu_n(dx) \rightarrow \int |x|^k \mu(dx)$ . The sequence  $(\mu_n)$  is called *uniformly integrable* if  $\sup_n \int |x| \mu_n(dx) 1(|x| > K) \rightarrow 0$  as  $K \rightarrow \infty$ . Let  $X, X_1, X_2, \dots$  be real-valued random variables. We say that  $X_n \rightarrow X$  weakly (resp. with weakly with  $k$ -th moments) if the corresponding probability distributions converge weakly (resp. weakly with  $k$ -th moments). We say that  $(X_n)$  is uniformly integrable if the collection of corresponding probability distributions is uniformly integrable.

**Lemma A.2.1.** *Let  $X_n$  and  $X$  be random numbers such that  $X_n \rightarrow X$  weakly with 1st moments. Then the sequence  $(X_n)$  is uniformly integrable.*

*Proof.* Given  $\epsilon > 0$ , by Lebesgue's dominated convergence we may choose  $K > 0$  such that  $\mathbb{E}X1(X > K) \leq \epsilon/3$ . Then let  $\phi_K$  be a continuous bounded function such that  $\phi_K(x) = x$  for  $x \leq K$  and  $\phi_K = 0$  for  $x \geq K + 1$ . Then

$$x1(x \leq K) \leq \phi_K(x) \leq x1(x \leq K + 1),$$

so that

$$\begin{aligned} \mathbb{E}X_n1(X_n > K + 1) &= \mathbb{E}X_n - \mathbb{E}X_n1(X_n \leq K + 1) \\ &\leq \mathbb{E}X_n - \mathbb{E}\phi_K(X_n) \\ &= \mathbb{E}X_n - \mathbb{E}\phi_K(X) + \mathbb{E}\phi_K(X) - \mathbb{E}\phi_K(X_n) \\ &\leq \mathbb{E}X_n - \mathbb{E}X1(X \leq K) + \mathbb{E}\phi_K(X) - \mathbb{E}\phi_K(X_n) \\ &= \mathbb{E}X1(X > K) + \mathbb{E}X_n - \mathbb{E}X + \mathbb{E}\phi_K(X) - \mathbb{E}\phi_K(X_n) \\ &\leq \epsilon/3 + |\mathbb{E}X_n - \mathbb{E}X| + |\mathbb{E}\phi_K(X_n) - \mathbb{E}\phi_K(X)|. \end{aligned}$$

Then we may choose  $n_0$  so large that  $|\mathbb{E}X_n - \mathbb{E}X| \leq \epsilon/3$  and  $|\mathbb{E}\phi_K(X_n) - \mathbb{E}\phi_K(X)| \leq \epsilon/3$  for all  $n > n_0$ . Hence  $\mathbb{E}X_n1(X_n > K + 1) \leq \epsilon$  for all  $n > n_0$ . Furthermore, for every  $1 \leq m \leq n_0$  we may choose, again by Lebesgue's dominated convergence,  $K_m$  such that  $\mathbb{E}X_m1(X_m > K_m) \leq \epsilon$ . Now if we choose  $L = \max\{K + 1, K_1, \dots, K_{n_0}\}$ , it follows that  $\sup_n \mathbb{E}X_n1(X_n > L) \leq \epsilon$ .  $\square$

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