

Network Traffic Measurements and Analysis

Lecture IV: Stochastic processes in network measurements

Markku Liinaharja (slides originally made by Esa Hyytiä),

Department of Communications and Networking Aalto University, School of Electrical Engineering

Version 0.2, October 4, 2017

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- Stochastic processes
- Self-similarity



Stochastic process - definition

- Mapping from the sample space to a function
 - Each random outcome corresponds to a realization of a stochastic process that is a function of "time"
- Sequence of random variables
 - ▶ Continuous-time processes $\{X_t, t \ge 0\}$
 - ▶ Discrete-time processes $\{X_n, n = 1, 2, ...\}$
- Defined by the n:th order distributions (for all n), e.g.,
 - ▶ First-order: $\{f_{X_1}(x), f_{X_2}(x), \ldots\}$
 - ► Second-order: $\{f_{X_1,X_2}(x), f_{X_1,X_3}(x), \ldots\}$
- All the distributions are needed to include all possible dependencies

Stationary processes

- Stochastic process is (strictly) stationary if all of its distribution functions are invariant under time shifts
- Simpler and more useful conditions, e.g.,
 - Wide-sense/weakly stationary process: mean and autocovariance are invariant in time
- 1st order statistics, mean
 - Expectation at time t, E(X_t)
- 2nd order statistics, autocovariance

$$R_{t,s} = \mathsf{E}((X_t - E[X_t])(X_s - E[X_s])) = \mathsf{Cov}[X_t, X_s].$$

▶ Wide-sense discrete-time stationary process, for all n, k:

$$E(X_n) = E(X_1), \quad Cov[X_n, X_{n+k}] = Cov[X_1, X_{k+1}].$$



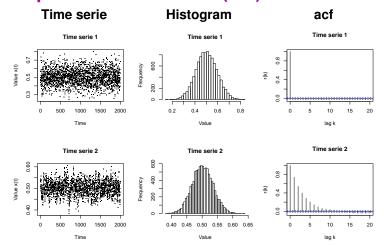
Autocorrelation

- If a process is at least wide-sense stationary, we can specify useful properties without regard to the particular time instant n
- Autocovariance of a wide-sense stationary process depends only on the lag k between random variables X_i and X_{i+k}
- Autocorrelation is normalized autocovariance

$$r(k) = \frac{\operatorname{Cov}[X_n, X_{n+k}]}{\operatorname{Cov}[X_n, X_n]}.$$



Example: autocorrelation (acf)



Samples in time serie 1 appear to be i.i.d. (no dependence)



Some stochastic processes of interest

Arrival process

- ► $\{A_n, n = 0, 1, ...\}$, where A_n is the time instant of the nth arrival
- Non-decreasing, non-stationary

Inter-arrival process

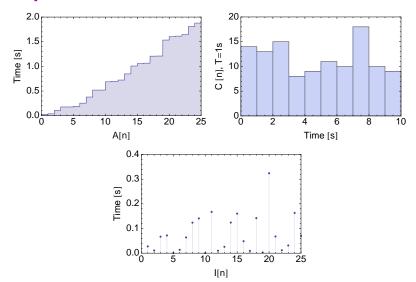
• $\{I_n, n = 1, 2, ...\}$, where $I_n = A_n - A_{n-1}$ is the length of the nth inter-arrival time

Time series of counts

- For a fixed time interval T
- ► Time series of counts $\{C_n, n = 0, 1, ...\}$, where $C_n = \#\{A_m | nT < A_m < (n+1)T\}$
- Arrivals can be packets, bits, bytes, ...
- Most common form of reporting network traffic



Example: session start times





Poisson process

- Inter-arrival process of Poisson process is such that
 - ▶ I_n and I_k are **independent** for $n \neq k$
 - ▶ I_n obeys Exp(λ) distribution for some parameter λ
- Alternatively, but equivalently: Poisson process is such that for any T > 0
 - ▶ C_n and C_k are independent for all $n \neq k$
 - $C_n \sim \text{Poisson}(\lambda T)$
- Poisson process is a widely used arrival process in performance modeling and analysis
 - "Infinitely" large population from which arrivals come independently
 - Often a valid assumption for e.g. human behavior in a large population
- Good for session arrivals (and phone calls ...)



Example

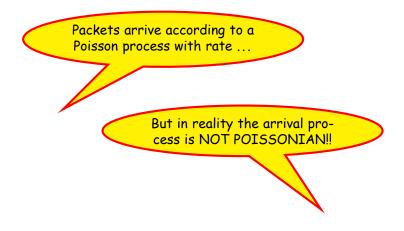
- Are the session start times from a Poisson process?
- Independent and exponentially distributed inter-arrivals?
- ACF, qqplots, ecdf, ...

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```
> arrivals<-scan("testidata.txt");
Read 4360 items
> fitdistr(arrivals,"exponential");
   rate
11.2151101
( 0.1698479)
> qqplot(arrivals, qexp(ppoints(arrivals), 11.21511));
abline(0,1, col = 2, lty = 2)
```



Poisson conversation ...





Dependence structure

- As a single flow can contribute packet or byte arrivals to several consecutive measurement periods, the independence assumption of Poisson process breaks down, e.g., for packet count processes
- If ACF plot reveals that past values contribute to the present other models must be considered
 - Note again that stationary processes are only good for describing stable data . . .
 - Network traffic is generally not stable at long time scales



Short- and Long-Range Dependence

1. Short-range dependence (SRD):

The coupling between values at different times decreases rapidly as the lag k increases

No tail: r(k) = 0 for $k > \theta$, or

Exp-tail: $r(k) \sim \beta^{-k}$ for $k > \theta$, (for some $\theta > 0$ and $\beta > 1$)

2. Long-range dependence (LRD)

The coupling between values at different times decreases slower than exponentially

Power-law tail: $r(k) \sim k^{-\beta}$, for $k > \theta$ (for some $\theta > 0$ and $0 < \beta < 1$)

SRD:
$$\sum_{k=1}^{\infty} r(k) < \infty$$
 LRD: $\sum_{k=1}^{\infty} r(k) = \infty$



Time series modeling

- ► Time series analysis provides suitable models for situations where autocorrelation shows significant dependence
- Examples from models
 - Moving average (MA)

$$Y_n = \mu + X_n + a_1 X_{n-1} + a_2 X_{n-2} + \ldots + a_k X_{n-k}$$
 where $E(X_n) = 0$ and $E(X_n^2) = \sigma^2$.

Autoregressive (AR)

$$Y_n = X_n + a_0 + a_1 Y_{n-1} + \ldots + a_k Y_{n-k}$$
 where $E(X_n) = 0$, $E(X_n^2) = \sigma^2$, and $E(X_n X_m) = 0$ for $n \neq m$.

- ► EWMA, ARMA, ARIMA, DAR, ...
- Well-developed field and especially widely used in signal processing, econometrics, etc.
- ► Cf., http://www.statsoft.com/textbook/sttimser.html



Contents

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- Self-similarity

Scaling

Long-range dependence is closely linked to a special property of network traffic; its **scaling behavior**

Rescaled view of traffic (e.g. time series of byte counts)

$$X_n^{(m)} \triangleq \sum_{i=nm}^{nm+m-1} X_i.$$

If the original data was observed on a time scale T, the rescaled process is observed on time scale Tm

Self-similarity

▶ A zero mean stochastic process $\{X_n\}$ is called **self-similar** with **Hurst parameter H** if, for all m, the aggregated process $\{X^{(m)}\}$ has the same distribution as $\{m^H X_n\}$, i.e.,

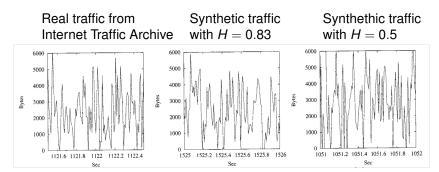
$$\{m^H X_n\} \stackrel{d}{=} \{X^{(m)}\}, \quad m > 0 \text{ and } 1/2 \le H < 1.$$

- Process can be self-similar only if it is LRD:
 - ▶ In fact, asymptotically autocorrelation $r(k) \sim k^{2H-2}$
 - ► Self-similarity means that $r^{(m)}(k) = r(k)$, with k, m > 0

"Self-similar processes have the same autocorrelation function on all time scales"

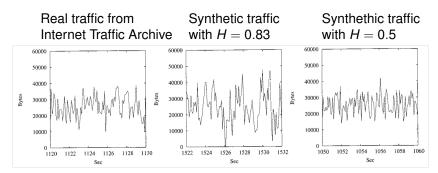
- (Nonself-similar: $r^{(m)}(k) \to 0$ as $m \to \infty$ for k = 0, 1, 2, ...)
- Self-similarity describes how the variability of the process scales





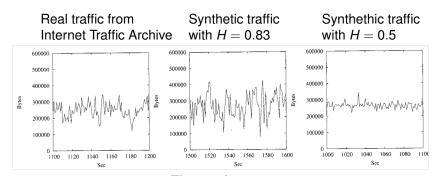
Timescale: 10 ms





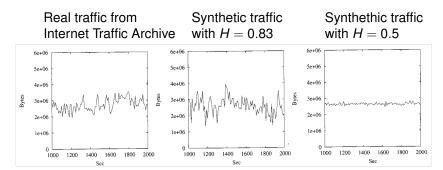
Timescale: 100 ms





Timescale: 1 s





Timescale: 10 s



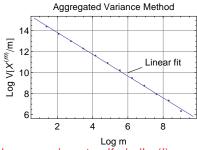
Measuring self-similarity

- Self-similarity results from variety of factors, but the main cause is the long-range dependency caused by heavytailed file/flow-size distributions
- Simplest way of estimating the Hurst parameter H is called the aggregated variance method:
 - Let X_n be a series of counts
 - ▶ Plot the variance of $m^{-1}X^{(m)}$ against m on a log-log scale
 - ▶ If the original data is well modeled with a self-similar process, then the variance of $m^{-1}X^{(m)}$ will follow m^{2H-2} Plot should show a straight line with slope $-\beta$ greater than -1
 - ► $H = 1 \beta/2$
- Other, more efficient, methods exist but are beyond the scope of our course



Example (simple)

- ▶ Let X_n be i.i.d. random variables, $X_n \sim \text{Pareto}(100, 1.5)$
- ▶ I.e., infinite variance, with a finite mean E(X) = 300
- ► Compute aggregates $X^{(m)}/m$ for $m = 2^1, 2^2, 2^3, \dots$ and their sample variance $s^2_{(m)}$
- Mean remains the same, E(X) = 300
- ▶ Plot log m against log $s_{(m)}^2$



Slope is -1, i.e., H = 0.5 and process is not self-similar(!)



Why self-similarity?

- Self-similarity is a parsimonious model to describe the variability of traffic on different time scales
- Self-similarity has fundamental effects on queuing behavior of traffic and hence it is interesting in performance analysis
- Recognition of the phenomenon is seen as a fundamental step forward in the way analysts think about traffic
- However, it is easy to mix with non-stationarity
 - Hurst parameter estimates from a non-stable data will be misleading



Literature

Mark Crovella, Balachander Krishnamurthy, Internet
 Measurement: Infrastructure, Traffic & Applications,
 John Wiley and Sons Ltd, 2006.

