# ECON-C4100 - Capstone: Econometrics I Lecture 4: Univariate regression

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- We postpone further discussion of regression level diagnostics to the lectures on multivariate regression.
- The reason for this is that they (adjusted R<sup>2</sup>, F-test), are more meaningful in the multivariate context.

- Economic interpretation & significance is of key importance.
- What about statistical significance?
- Recall the discussion on the properties of the sample mean:
  - It is a random variable (every random sample produces its own mean to be used as an estimate of the population mean).
  - 2 It is unbiased and consistent
  - **3** It has a distribution that we can characterize (with large *n*, becomes / approaches a normal distribution).

- Similarly, the parameters we "find" or estimate with OLS depend on the random sample available to us.
- Were you to draw a different random sample, you would get different parameter estimates.
- In other words, *beta*<sub>0</sub> and *beta*<sub>1</sub> are also random variables.
- We'd like to know their properties, i.e., how they are distributed, and how different things affect that distribution.
- Under assumptions that we'll discuss in a moment,  $\beta_0$  and  $\beta_1$  are (bivariate) normally distributed with a known mean and variance.

- $\hat{eta_0}$  and  $\hat{eta_1}$  are
  - unbiased
  - 2 consistent and
  - **3** efficient (with an extra assumption).
- under a set of assumptions.

- One needs to understand the assumptions that allow a particular interpretation of the results.
- Crucial to understand the assumptions & their implications.
- Crucial to form an opinion or test the validity of assumptions and/or the robustness of results to those assumptions.

- **1** Strict exogeneity:  $\mathbb{E}(u_i|X_i) = 0$ .
- (X<sub>i</sub>, Y<sub>i</sub>), i = 1, ..., n are independent and identically distributed across observations.
- **3**  $X_i$  and  $Y_i$  have finite *fourth* moments.
- 4 Auxiliary:  $u_i$  is homoscedastic.

$$E(u_i|X_i)=0$$

- Implies that *u* and *X* are uncorrelated.
- $E(u_i|X_i) = 0 \implies cov(u, X) = 0.$
- Not the other way round because correlation is about a linear relationship only.

- (X<sub>i</sub>, Y<sub>i</sub>), i = 1, ..., n are i.i.d.
- The same concept as before, but now over a joint distribution of two variables.
- Experiments where X chosen.
- Time series.

- $X_i$  and  $Y_i$  have finite *fourth* moments:  $\mathbb{E}(X)^4$ ,  $\mathbb{E}(Y)^4 < \infty$ .
  - = they have finite kurtosis.
- Needed to ensure that the standard errors are from a normal distribution (4th moment  $\approx$  variance of variance).
- Means that large outliers are (extremely) unlikely.

- $var(u_i|X_i = x) = \sigma^2$  for i = 1, ..., n.
- *u<sub>i</sub>* is homoscedastic (as opposed to heteroscedastic).
- Alternative:  $var(u_i|X_i = x) = \sigma_i^2$ .

Gauss Markov:

If A.1 - A.4 hold, then OLS is BLUE (Best Linear Conditionally Unbiased Estimator).

• In practice, data have/lead to heteroscedastic errors almost always.

 $\rightarrow$  easy and efficient ways to correct for heteroscedasticity.

- Modern default is to use (heteroscedasticity) robust standard errors.
- Wrong assumption on variance of the error term biases standard errors, *not coefficients*.

$$Y_i = \beta_0 + \beta_1 X_i + u = \mathbf{X}'_i \boldsymbol{\beta} + u_i$$
<sup>(1)</sup>

- Let's vary different aspects of the Data Generating Process (GDP).
- How do we do this?

1-1

- Let's use artificial data that has "appealing" features.
- Artificial data: ask the computer to generate it.

 $\rightarrow$  the researcher chooses what the data looks like.

- Monte Carlo simulation: repeat a statistical model S times on artificial data, look at means and distributions of parameters.
- We are going to generate artificial data that has the key properties of our FLEED data and use it to illustrate the effects of the OLS assumptions.

- 1 Decide the properties you want the data to have.
- 2 Choose the parameters of the model.
- 3 Use a random number generator to generate the exogenous variables, including the error term.
- Generate the dependent variable using the parameters and the exogenous variables.

#### Stata code

```
regr income age if year == 15 & income != .
2
  predict u_hat, res
3
  sum income age u_hat
  matrix beta
                 = e(b)
4
5
  matrix list beta
6
  scalar beta0 = beta[1,2]
7
  scalar beta1 = beta[1,1]
8
  qui sum u_hat if e(sample)
9
  scalar u_sd
               = r (sd)
10
  scalar age_m = r(mean)
11 scalar age_sd = r(sd)
```

#### Stata code

```
dropp all
 2
   global age_m
                       = age_m
 3
   global b0
                       = beta0
4
   global b1
                       = beta1
5
   global age_m
                        = age_m
6
   set seed 987345
7
8
   capture program drop myprog_sim
9
   program define myprog_sim
10
     drop _all
11
     set obs 10000
12
                          = $age_m + rnormal(0,age_sd)
     gen x
13
                             = $h0
     scalar beta0
14
     scalar betal
                              = $b1
15
                       = rnormal(0, u_sd)
     gen u
16
     aui sum u
17
     scalar u₋mean
                        = r(mean)
18
     replace u
                        = u — u mean
19
                            = beta0 + beta1 * x + u
     gen y
20
     regr v x
21
   end
22
23
   simulate _b _se, saving(myprog_sim, replace) reps(1000); myprog_sim
24
   display "OLS nobs 10000"
25
   sum
```

- Increasing sample size
  - **1** Brings the coefficients closer to their true value.
  - 2 Reduces the standard errors of the coefficients.

- The data generating process:
- Case #1:  $u = rnormal(0, \sigma_u^2)$
- Case #2:  $u_{het} = rnormal(0, \sigma_u^2) \times (1 + z \times age)$ 
  - z a multiplier chosen by the modeller.
- Notice both cases satisfy  $\mathbb{E}(u|age) = 0$ .

$$Income_i = \hat{\beta}_0 + \hat{\beta}_1 age_i + u_i$$

$$Income_{het,i} = \hat{\beta_0} + \hat{\beta_1}age_i + u_{het,i}$$

- Let's vary *z* = 1, ..., 10
- Sample size 1000.

. estimates store lin\_est

. estimates table lin\_est, b(%7.3f) se(%7.3f) p(%7.3f) stat(r2

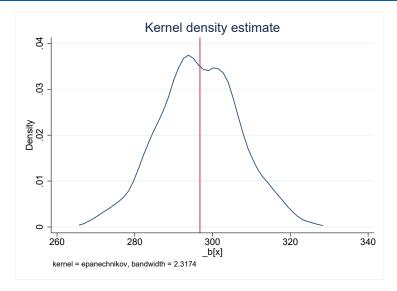
Variable	lin_est
age _cons	296.754 13.353 0.000 1.1e+04 607.567 0.000
r2	0.076
	1 1 / /

legend: b/se/p

## Let's first check how this works with z = 0: Estimates

Variable	Obs	Mean	Std. Dev.	Min	Max
_b_x	1,000	296.5223	10.36742	267.9936	325.9853
_b_cons	1,000	10664.63	441.6165	9412.67	11876.38
_se_x	1,000	10.31445	.1038239	9.895374	10.73409
_se_cons	1,000	469.3358	4.696664	449.9725	488.3702

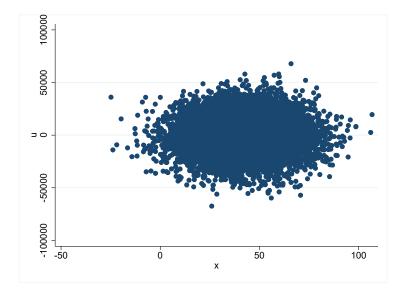
# Let's first check how this works with z = 0 Distribution of $\beta_1$



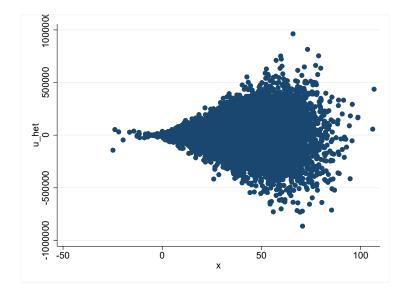
round	$\beta_1$	$se_{\beta_1}$	$\beta_0$	$se_{eta_0}$
1	293.58	174.12	10741.91	7815.36
2	337.49	318.82	9111.05	14304.02
3	297.95	462.01	10402.12	20770.64
4	261.93	606.49	11762.07	27183.05
5	326.50	747.35	9278.87	33530.71
6	374.37	895.85	7290.83	40201.69
7	176.23	1042.51	15146.60	46913.89
8	397.26	1186.82	6880.53	53272.23
9	320.64	1328.07	9650.81	59678.00
10	358.15	1470.39	8538.93	66067.65

- $\mathbb{E}[u|X] = 0$  holds for all the samples.
- But the variance of *u* becomes an increasing function of *age*.
- This leads to a very different looking distribution.

## Distribution of homoscedastic u



### Distribution of heteroscedastic *u*



- (large) outliers may lead to biased estimates.
- Difficulty is of course to determine what is large.
- For illustration, let's replace a few values of *age* with much larger values.
- First, *age* of one individual multiplied by 10 ("typo") in a sample of 1000 observations.
- Second, same done for 10 individuals.

% obs. changed	$\beta_1$	$se_{\beta_1}$	$\beta_0$	$se_{eta_0}$
0.1	47.29	12.71	20958.43	795.41
1	207.16	26.63	14348.99	1247.97