

# ECON-C4100 - Capstone: Econometrics I

## Lecture 6: Multiple regression #1: estimation

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- At the end of this lectures, you
  - 1 understand what how multivariate regression differs from univariate regression.
  - 2 understand how and why to carry out a multivariate regression analysis.
  - 3 appreciate the assumptions made in multivariate regression analysis.
  - 4 are aware of the most common pitfalls in regression analysis.

## Starting point:

$$Y = f(X_1, X_2, \dots, X_k, u)$$

Outcome variable of interest a function of several variables. Observables and unobservables. One or more hypotheses (needed)?

- ① Is income affected by age?
  - ② Do women and men of same age earn differently?
- Let's study these using the open access FLEED data of Statistics Finland.
  - These data can be downloaded from the Statistics Finland [web page](#).
  - We will use the year 15 (= 2010) cross section data.

$$Y = f(X, u) = \beta_0 + \beta_1 X + U$$

$$\mathbb{E}[Y|X = x] = \beta_0 + \beta_1 x$$

# Multivariate/multiple regression

$$Y = f(X_1, X_2, u) = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + u$$

$$E[Y|\mathbf{X} = \mathbf{x}] = \beta_0 + \beta_1 x_1 + \beta_2 x_2$$

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2$$

- This is the so called **population regression line (populaatio regressio)**.
- $Y =$  **dependent variable (vastemuuttuja)** or **endogenous variable**.
- $X_k =$  **independent variable  $k$**  (selittävä muuttuja) or **exogenous variable  $k$**  or *regressor  $k$* .
- $\beta_0, \beta_1, \beta_2$ : **parameters of the model**.

# Are all Xs born equal?

- Depends...
- Treatment variable = the one of primary interest.
- Control variable(s) = affect(s)  $Y$  but we are not (so much) interested in this/these.
- Why include variables that we are not interested in?



# What type of control variables matter, how & why?

- ①  $cov(X_1, X_2) = 0$
  - ②  $cov(X_1, X_2) \neq 0$
- Key is whether the treatment variable and control variable are correlated or not.

# What type of control variables matter, how & why?

- Why is this key? Recall

$$\hat{\beta}_1 = \beta_1 + \rho_{xu} \frac{\sigma_u}{\sigma_x}. \quad (1)$$

Rewrite

$$u = \beta_2 X_2 + v \quad (2)$$

# What type of control variables matter, how & why?

Assume

$$\text{cov}(\mathbf{X}, v) = 0$$

Then

$$\hat{\beta}_1 = \beta_1 + \beta_2 \rho_{X_1 X_2} \times \frac{\sigma_{X_2}}{\sigma_{X_1}} \quad (3)$$

Make sure you know how to derive equation (3).

# What type of control variables matter, how & why?

If the  $X$ s are correlated, then the bias in  $\beta_1$  depends on

- 1 the impact of  $X_2$  on  $Y$ .
- 2 the correlation between the  $X$ s.
- 3 how much variance  $X_2$  has relative to  $X_1$ .

# What type of control variables matter, how & why?

- So are we home if  $\text{cov}(X_1, X_2) = 0$ ?
- Yes and no.
- If  $\text{cov}(X_1, X_2) = 0$  then  $\hat{\beta}_1 = \beta_1$
- However, adding  $X_2$  decreases the standard error / increases the precision of  $\hat{\beta}_1$ .

# What type of control variables matter, how & why?

- A two-variable model (App 6.2. in S&W).
- 2 explanatory variables and homosced. errors,  $\rho_{X_1, X_2} = 0$ . Then

$$\sigma_{\beta_1}^2 = \frac{1}{n} \frac{\sigma_u^2}{\sigma_{X_1}^2}$$

which is the variance of  $\hat{\beta}_1$ .

- Adding  $X_2$  necessarily decreases  $\sigma_u^2$ .

# Income modeled as function of age and gender?

- Let's look at the following model:

$$Income_i = \beta_0 + \beta_{AgeMV} Age_i + \beta_{GMV} G_i + u_{MV_i}$$

Where

$Age_i$  = age in years.

$G_i$  = dummy for gender.

$MV$  stands for **M**ultivariate.

# Should we suspect that age affects income?

- Experience increases with age.
- In a cross-section such as ours, younger people typically better educated than older (conditional on not being too young).
- Physical condition and mental agility start to decrease relatively early.



# Should we suspect that gender affects income?

- Segregation of job market a well known phenomenon.
- Women bear a larger share of household work and stay longer at home after getting a child.
- Educational levels and fields differ by gender.

# Some conditional descriptive statistics

```
. tabstat income age if year == 15, stat(mean sd p50) by(gender)
```

Summary statistics: mean, sd, p50  
by categories of: gender

gender	income	age
0	25478.2	41.5928
	18894.06	16.10072
	24000	42
1	21053.65	42.15153
	14852.83	16.48018
	20000	43
Total	23296.67	41.86563
	17163.61	16.28821
	21000	43

# Are age and gender correlated in our data?

```
. pwcorr age gender if year == 15, sig
```

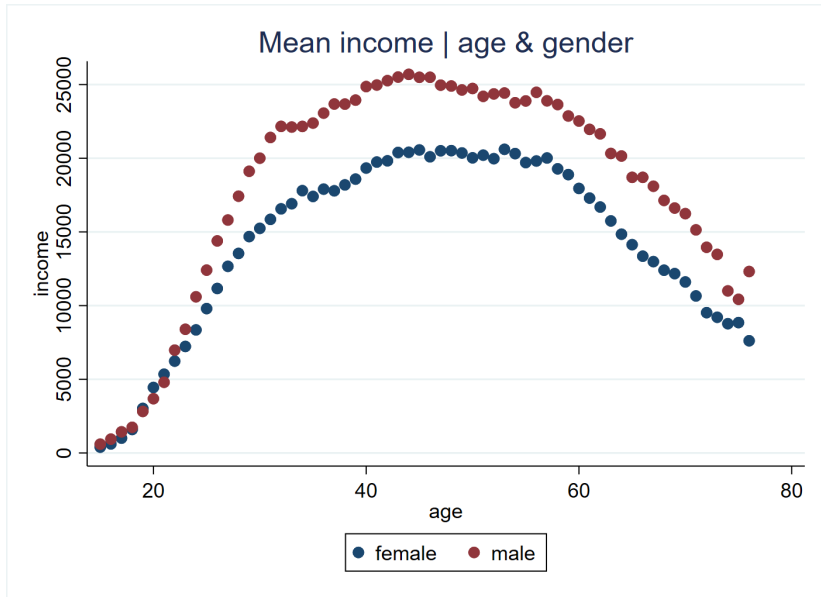
	age	gender
age	1.0000	
gender	0.0171 0.1755	1.0000

# Mean income conditional on age and gender

## Stata code

```
1 bysort age gender: egen income_m_age-g = mean(income)
2 bysort age gender: gen win_age-g_ind = _n
3 twoway scatter income_m_age-g age if gender == 1 & win_age-g_ind == 1 || ///
4   scatter income_m_age-g age if gender == 0 & win_age-g_ind == 1, ///
5   legend(lab (1 "female") lab (2 "male")) ///
6   graphregion(fcolor(white)) ///
7   ytitle("income") ///
8   title("Mean income | age & gender")
9 graph export "mean_income_age-gender.png", replace
```

# Mean income conditional on age and gender

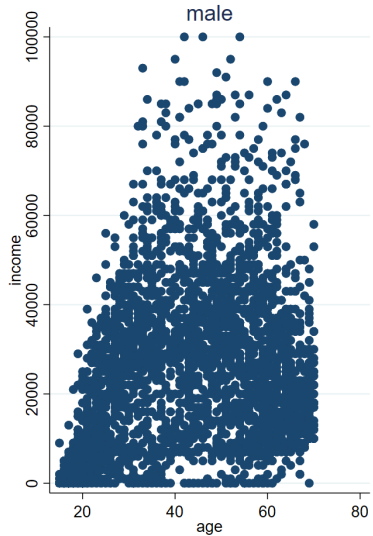
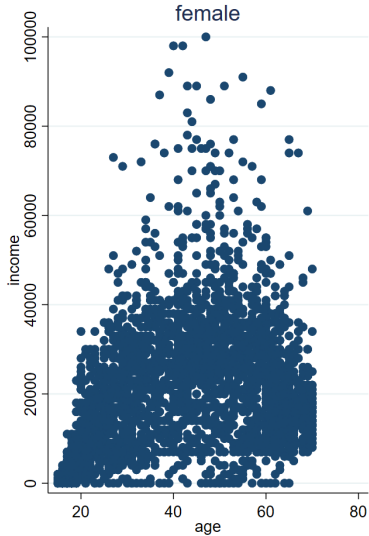


# Income-age scatter by gender

## Stata code

```
1 scatter income age if gender == 0 & year == 15, ///
2   graphregion(fcolor(white)) ///
3   title("male") ///
4   saving(income_age_male, replace)
5 scatter income age if gender == 1 & year == 15, ///
6   graphregion(fcolor(white)) ///
7   title("female") ///
8   saving(income_age_female, replace)
9 gr combine income_age_female.gph income_age_male.gph
10 graph export "income_age_gender.png", replace
```

# Income-age scatter by gender



# How to get $\beta_0$ , $\beta_1$ & $\beta_2$ : OLS

$$\min_{\beta_0, \beta_1, \beta_2} \sum_{i=1}^n [Y_i - (\beta_0 + \beta_1 X_{1i} + \beta_2 X_{2i})]^2 \quad (4)$$

$$\min_{\beta} (\mathbf{Y} - \mathbf{X}\beta)'(\mathbf{Y} - \mathbf{X}\beta) \quad (5)$$



## How to get $\beta_0$ , $\beta_1$ & $\beta_2$ : OLS

$$\hat{\beta}_1 = \frac{\sum x_2^2 \sum x_2 y - \sum x_1 x_2 \sum x_1 y}{\sum x_1^2 \sum x_2^2 - (\sum x_1 x_2)^2}$$

$$\hat{\beta} = (\mathbf{X}'\mathbf{X})^{-1}(\mathbf{X}'\mathbf{Y})$$

- Note: now not using matrix algebra leads to very cumbersome mathematics; with matrix algebra, the solution stays the same as with univariate regression.
- The expression for  $\hat{\beta}_2$  is symmetric with that of  $\hat{\beta}_1$ .
- Finally,  $\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}_1 - \hat{\beta}_2 \bar{X}_2$ .

# Let's compare univariate regressions to multivariate regression

$$Income = \beta_{0AgeUV} + \beta_{AgeUV} Age + u_{AgeUV} \quad (6)$$

$$Income = \beta_{0GUV} + \beta_{GUV} G + u_{GUV} \quad (7)$$

$$Income = \beta_0 + \beta_{AgeMV} Age + \beta_{GMV} G + u_{MV} \quad (8)$$

**Table:** Univariate income regressions

	(1)	(2)
	income	income
Age	296.8*** (13.35)	
Gender		-4424.6*** (440.5)
Constant	10654.7*** (607.6)	25478.2*** (309.3)
Observations	5973	5973
r <sup>2</sup>	0.0764	0.0166
F	493.9	100.9

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$

# Multivariate regression

**Table:** Income regressions

	(1)	(2)	(3)
	income	income	income
Age	296.8*** (13.35)		298.5*** (13.23)
Gender		-4424.6*** (440.5)	-4545.0*** (422.9)
Constant	10654.7*** (607.6)	25478.2*** (309.3)	12819.2*** (634.6)
Observations	5973	5973	5973
r2	0.0764	0.0166	0.0939
F	493.9	100.9	309.4

Standard errors in parentheses

\*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$

- ① How do the individual coefficients compare to univariate results?
- ② What explains the difference(s)?
- ③ What about statistical significance of individual coefficients?
- ④ What about several / all coefficients?
- ⑤ What about  $R^2$ ?

- 6 What is the interpretation of individual coefficients?
- 7 (under what assumptions) does OLS work?
- 8 How to choose which explanatory variables to include / exclude?
- 9 What if the world is more complicated than linear?
- 10 What all can go wrong, and how would I know / find out?

# Q1 & Q2 multivariate vs. univariate?

- 1 How do the individual coefficients compare to univariate results?
- 2 What explains the difference(s)?

## Q1 & Q2 multivariate vs. univariate?

- Compare the *Age* coefficient in the univariate to that in the multivariate regression.
- What can you conclude? Recall

$$\hat{\beta}_{AgeUV} = \beta_{AgeMV} + \beta_{GMV} \rho_{Age,G} \times \frac{\sigma_G}{\sigma_{Age}}$$



# Multivariate regression

```
. estout income_age income_gender income_age_gender, cells(b(star fmt(3)) se(par fmt(3)))
```

	income_age b/se	income_gen~r b/se	income_age~r b/se
age	296.754*** (13.353)		298.548*** (13.228)
gender		-4424.553*** (440.537)	-4545.022*** (422.935)
_cons	10654.695*** (607.567)	25478.203*** (309.335)	12819.203*** (634.636)

```
. pwcorr age gender if e(sample)
```

	age	gender
age	1.0000	
gender	0.0126	1.0000

```
. sum age gender if e(sample)
```

Variable	Obs	Mean	Std. Dev.	Min	Max
age	5,973	42.60087	15.9866	15	70
gender	5,973	.4930521	.4999936	0	1

## Q1 & Q2 multivariate vs. univariate?

- Plug numbers from the previous slide into the bias formula:

$$\begin{aligned} \text{Bias}_{\beta_{\text{ageUV}}} &= \beta_{\text{GMV}} \rho_{\text{Age},G} \times \frac{\sigma_G}{\sigma_{\text{Age}}} \\ &= -4545.02 \times 0.0126 \times \frac{0.500}{15.987} = -1.791 \end{aligned}$$

- Compare to

$$\beta_{\text{AgeUV}} - \beta_{\text{AgeMV}} = 296.754 - 298.548 = -1.794$$

- Do the same for gender.
- What can you conclude?

# Q1 & Q2 multivariate vs. univariate?

- Multivariate regression allows the researcher to
  - ① control for observable variables and thereby either remove (omitted variable) bias and/or increase efficiency.
  - ② test several hypotheses simultaneously.
  - ③ (as we will see), enrich the main hypotheses to allow for heterogeneous effects.

- 3 What about statistical significance of individual coefficients?
- 4 What about the statistical significance of several / all coefficients?

- Can we reject the null that
  - ①  $\beta_0 = 0$ ,
  - ②  $\beta_{Age} = 0$ ,
  - ③  $\beta_G = 0$ ?

# Q3 & Q4 statistical significance, individual coefficients

```
. regr income age gender          if year == 15
```

Source	SS	df	MS	Number of obs	=	5,973
Model	1.6524e+11	2	8.2622e+10	F(2, 5970)	=	309.43
Residual	1.5940e+12	5,970	267009188	Prob > F	=	0.0000
				R-squared	=	0.0939
				Adj R-squared	=	0.0936
Total	1.7593e+12	5,972	294589468	Root MSE	=	16340

income	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
age	298.5478	13.22762	22.57	0.000	272.6169	324.4787
gender	-4545.022	422.9347	-10.75	0.000	-5374.127	-3715.917
_cons	12819.2	634.6356	20.20	0.000	11575.09	14063.32

- Are  $\beta_0, \beta_{Age}, \beta_g$  all = 0?
- F - test (and others) for the *joint significance*.
- Cannot do this by looking at individual (t-) tests.
- Reason: two or more random variables  $\rightarrow$  need their joint distribution.

- F test (under homoscedasticity). For illustration only.

$$F = \frac{(SSR_{restricted} - SSR_{unrestricted})/q}{SSR_{unrestricted}/(n - k_{restricted} - 1)}$$
$$= \frac{(R^2_{restricted} - R^2_{unrestricted})/q}{(1 - R^2_{unrestricted})/(n - k_{restricted} - 1)}$$

- Modern software calculate the heteroscedasticity-robust F-test.



# Q3 & Q4 statistical significance, individual coefficients

```
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```

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- What about  $\beta_{Age} = \beta_G = 0$  ?
- In other words, Null hypothesis is that a subset of parameters are zero.
- Modern software allow this.

# Q3 & Q4 statistical significance, individual coefficients

```
. testparm age gender  
  
( 1) age = 0  
( 2) gender = 0  
  
F( 2, 5970) = 309.43  
Prob > F = 0.0000
```

- What about  $\beta_{Age} = \beta_G$ ?
- Need either a direct test modern software allow this (easily).
- Or a trick (add and subtract).

## Q3 & Q4 statistical significance, individual coefficients

```
. test age = gender  
  
( 1) age - gender = 0  
  
F( 1, 5970) = 130.92  
Prob > F = 0.0000
```

- With multivariate regression:
  - ① Important to check the regression diagnostic statistics (F-test) (more on this to follow).
  - ② Rich possibilities to test hypotheses that involve multiple parameters.

## Q5 What about $R^2$ ?

$$R^2 = \frac{ESS}{TSS} = 1 - \frac{SSR}{TSS}$$

- $R^2$  increases (almost) surely as you add explanatory variables.
- Adjusted  $R^2$  corrects for this:

$$adjR^2 = 1 - \frac{n-1}{n-k-1} \frac{SSR}{TSS} = \frac{s_{\hat{u}}^2}{s_Y^2}$$

- $n$  = number of obs;  $k$  = number of expl. variables.
- Adjusted  $R^2$  always lower than  $R^2$ .

# Q5 What about $R^2$ ?

```
. regr income age gender          if year == 15
```

Source	SS	df	MS	Number of obs	=	5,973
Model	1.6524e+11	2	8.2622e+10	F(2, 5970)	=	309.43
Residual	1.5940e+12	5,970	267009188	Prob > F	=	0.0000
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## Q5 What about $R^2$ ?

- High  $R^2$  / an increases in  $R^2$  says nothing about causality.
- High  $R^2$  does not mean your model does not suffer from omitted variable bias.
- High  $R^2$  does not mean you have the right set of explanatory variables.
- High  $R^2$  tells nothing about the economic significance of your results.
- High  $R^2$  means that factors outside your model (= the stuff going into the error term) play a relatively speaking smaller role in the process that determines the value of  $Y$ .
- But, as we saw from the F-test formula, (changes in)  $R^2$  are indicative and a certain level of  $R^2$  is needed to reject the Null that all your model parameters are insignificant.