Mathematics for Economists
Aalto BIZ
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Juuso Välimäki

## Economic Applications

## Ordinary least squares

Consider a statistical sample consisting of on $N$ pairs of observations

$$
\left(y_{1}, x_{1}\right), \ldots,\left(y_{N}, x_{N}\right)
$$

Suppose that we want to find a linear relation between $x$ and $y$. We would like to find a coefficient $\beta$ that rationalizes the observations as

$$
y_{i}=\beta x_{i} .
$$

If we have many observations, this will not be satisfied in general. To account for errors in the linear relationship, specify the following statistical model

$$
y_{i}=\beta x_{i}+\varepsilon_{i},
$$

where $\varepsilon_{i}$ is an identically and independently distributed error term for all $i$.

Our task is to infer $\beta$ from the sample. One way of doing this is based on minimizing the sum of squared error terms $\left(\varepsilon_{i}\right)^{2}$, i.e. to

$$
\min _{\beta} f(\beta)=\sum_{i=1}^{N}\left(y_{i}-\beta x_{i}\right)^{2}
$$

Compute $f^{\prime}(\beta)$ and consider $\widehat{\beta}$ such that:

$$
f^{\prime}(\widehat{\beta})=0
$$

By taking the derivative, we get:

$$
f^{\prime}(\widehat{\beta})=\sum_{i=1}^{N}-2 x_{i}\left(y_{i}-\widehat{\beta} x_{i}\right) .
$$

As a result, $f^{\prime}(\widehat{\beta})=0$ if

$$
\widehat{\beta}=\frac{\sum_{i=1}^{N} x_{i} y_{i}}{\sum_{i=1}^{N} x_{i}^{2}}
$$

Since $f^{\prime \prime}(\hat{\beta})=\sum_{i=1}^{N} x_{i}^{2}>0$, we have found the minimum.
If we want to include a constant term $\alpha$, we get:

$$
y_{i}=\alpha+\beta x_{i}+\varepsilon_{i} .
$$

function of $(\alpha, \beta)$ :

$$
f(\alpha, \beta)=\sum_{i=1}^{N}\left(y_{i}-\alpha-\beta x_{i}\right)^{2}
$$

To find $(\widehat{\alpha}, \widehat{\beta})$ such that

$$
\frac{\partial f(\widehat{\alpha}, \widehat{\beta})}{\partial \alpha}=\frac{\partial f(\widehat{\alpha}, \widehat{\beta})}{\partial \beta}=0
$$

we get:

$$
\begin{aligned}
\sum_{i=1}^{N}\left(y_{i}-\widehat{\alpha}-\widehat{\beta} x_{i}\right) & =0 \\
\Sigma_{i=1}^{N}-2 x_{i}\left(y_{i}-\widehat{\alpha}-\widehat{\beta} x_{i}\right) & =0 .
\end{aligned}
$$

Solving for $\alpha$ from the first equation gives:

$$
\widehat{\alpha}=\frac{\Sigma_{i=1}^{N} y_{i}-\widehat{\beta} \Sigma_{i=1}^{N} x_{i}}{N}:=\bar{y}-\widehat{\beta} \bar{x} .
$$

Using the first equation we also see that:

$$
\Sigma_{i=1}^{N} \bar{x}\left(y_{i}-\widehat{\alpha}-\widehat{\beta} x_{i}\right)=0
$$

By substituting into the second, we get:

$$
\widehat{\beta}=\frac{\sum_{i=1}^{N}\left(x_{i}-\bar{x}\right)\left(y_{i}-\bar{y}\right)}{\sum_{i=1}^{N}\left(x_{i}-\bar{x}\right)^{2}}=\frac{\operatorname{Cov}(y, x)}{\operatorname{Var}(x)} .
$$

More generally, we can consider samples with more explanatory variables: $\left(y_{1}, x_{11}, x_{21}, \ldots, x_{K 1}\right), \ldots\left(y_{N}, x_{1 N}, \ldots, x_{K N}\right)$ and a linear model

$$
\left(\begin{array}{c}
y_{1} \\
\vdots \\
y_{n}
\end{array}\right)\left(\begin{array}{c}
= \\
\vdots \\
=
\end{array}\right)\left(\begin{array}{ccc}
\beta_{1} x_{11}+ & \cdots & \beta_{K} x_{K 1} \\
\vdots & & \vdots \\
\beta_{1} x_{1 N} & \cdots & \beta_{K} x_{K N}
\end{array}\right)+\left(\begin{array}{c}
\varepsilon_{1} \\
\vdots \\
\varepsilon_{N}
\end{array}\right)
$$

or in matrix form:

$$
\boldsymbol{y}=\boldsymbol{X} \boldsymbol{\beta}+\boldsymbol{\varepsilon}
$$

We can compute the sum of squares now as:

$$
\begin{aligned}
f(\boldsymbol{\beta}) & =\varepsilon \cdot \varepsilon=(\boldsymbol{y}-\boldsymbol{X} \boldsymbol{\beta})^{\top}(\boldsymbol{y}-\boldsymbol{X} \boldsymbol{\beta}) \\
& =\boldsymbol{y} \cdot \boldsymbol{y}-(\boldsymbol{X} \boldsymbol{\beta})^{\top} \boldsymbol{y}-\boldsymbol{y}^{\top} \boldsymbol{X} \boldsymbol{\beta}+\boldsymbol{\beta}^{\top} \boldsymbol{X}^{\top} \boldsymbol{X} \boldsymbol{\beta} \\
& =\boldsymbol{y} \cdot \boldsymbol{y}-2 \boldsymbol{y}^{\top} \boldsymbol{X} \boldsymbol{\beta}+\boldsymbol{\beta}^{\top} \boldsymbol{X}^{\top} \boldsymbol{X} \boldsymbol{\beta} .
\end{aligned}
$$

The derivative of $-2 \boldsymbol{y}^{\top} \boldsymbol{X} \boldsymbol{\beta}$ is the row vector $-2 \boldsymbol{y}^{\top} \boldsymbol{X}$ and the derivative of $\boldsymbol{\beta}^{\top} \boldsymbol{X}^{\top} \boldsymbol{X} \boldsymbol{\beta}$ is $2 \boldsymbol{\beta}^{\top} \boldsymbol{X}^{\top} \boldsymbol{X}$. (To see this, write the matrix product as a sum). Writing with the gradient, we have:

$$
\nabla f(\widehat{\boldsymbol{\beta}})=-2 \boldsymbol{X}^{\top} \boldsymbol{y}+2 \boldsymbol{X}^{\top} \boldsymbol{X} \widehat{\boldsymbol{\beta}}
$$

Therefore we can find a candidate for the extremum by setting

$$
\nabla f(\widehat{\boldsymbol{\beta}})=0
$$

Solving for $\boldsymbol{\beta}$, we get:

$$
\widehat{\boldsymbol{\beta}}=\left(\boldsymbol{X}^{\top} \boldsymbol{X}\right)^{-1} \boldsymbol{X}^{\top} \boldsymbol{y}
$$

Exercise: Is there a need to add a constant term to get a more general formula? Show also that the Hessian of $f(\boldsymbol{\beta})$ is positive definite.

## Second derivatives of functions used in economics

## Power function

Sums of power functions are common in consumer theory. Each component in the sum measures the utility from consumption in a given period.
(Individual components $x_{i}^{\rho}$ are often called CRRA functions since they display constant relative risk aversion).

$$
f\left(x_{1}, x_{2}\right)=x_{1}^{\rho}+x_{2}^{\rho} .
$$

Form the gradient

$$
\nabla f\left(x_{1}, x_{2}\right)=\binom{\frac{\partial f\left(x_{1}, x_{2}\right)}{\partial x_{1}}}{\frac{\partial f\left(x_{1}, x_{2}\right)}{\partial x_{2}}}=\binom{\rho x_{1}^{\rho-1}}{\rho x_{2}^{\rho-1}} .
$$

Form the Hessian matrix by taking the derivative of the gradient:

$$
H f\left(x_{1}, x_{2}\right)=\left(\begin{array}{ll}
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} \partial x_{1}} & \frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} x_{2}} \\
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{1}} & \frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{2}}
\end{array}\right) .
$$

We get:

$$
H f\left(x_{1}, x_{2}\right)=\left(\begin{array}{cc}
\rho(\rho-1) x_{1}^{\rho-2} & 0 \\
0 & \rho(\rho-1) x_{2}^{\rho-2}
\end{array}\right)
$$

$D^{2} f\left(x_{1}, x_{2}\right)$ is thus negative definite when $x_{i} \neq 0$ ja $0<\rho<1$.

## CES -function

Recall the CES -function (utility and production function).

$$
f\left(x_{1}, x_{2}\right)=\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}} .
$$

Form the gradient:

$$
\nabla f\left(x_{1}, x_{2}\right)=\binom{\frac{\partial f\left(x_{1}, x_{2}\right)}{\partial x_{1}}}{\frac{\partial f\left(x_{1}, x_{2}\right)}{\partial x_{2}}}=\binom{\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-1} x_{1}^{\rho-1}}{\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-1} x_{2}^{\rho-1}} .
$$

The the Hessian matrix is:

$$
H f\left(x_{1}, x_{2}\right)=\left(\begin{array}{ll}
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} \partial x_{1}} & \frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} 1 x_{2}} \\
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{1}} & \frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{2}}
\end{array}\right) .
$$

By the product rule:

$$
\begin{aligned}
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} \partial x_{1}}= & (\rho-1) x_{1}^{\rho-2}\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-1} \\
& +\left(\frac{1}{\rho}-1\right)\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-2} \rho x_{1}^{2 \rho-2}
\end{aligned}
$$

$$
\begin{aligned}
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} \partial x_{2}}= & \left(\frac{1}{\rho}-1\right)\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-2} \rho x_{2}^{\rho-1} x_{1}^{\rho-1} \\
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{2}}= & (\rho-1) x_{2}^{\rho-2}\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-1} \\
& +\left(\frac{1}{\rho}-1\right)\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-2} \rho x_{2}^{2 \rho-2} .
\end{aligned}
$$

By collecting the common terms, we get:

$$
\begin{gathered}
D^{2} f\left(x_{1}, x_{2}\right)=\left(\begin{array}{cc}
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} \partial x_{1}} & \frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{1} 1 x_{2}} \\
\frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{1}} & \frac{\partial^{2} f\left(x_{1}, x_{2}\right)}{\partial x_{2} \partial x_{2}}
\end{array}\right) \\
=\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-2}\left(\begin{array}{cc}
(\rho-1) x_{1}^{\rho-2} x_{2}^{\rho} & (1-\rho) x_{2}^{\rho-1} x_{1}^{\rho-1} \\
(1-\rho) x_{2}^{\rho-1} x_{1}^{\rho-1} & (\rho-1) x_{2}^{\rho-2} x_{1}^{\rho}
\end{array}\right) .
\end{gathered}
$$

When computing the determinant, we can separate the common factor:

$$
\begin{gathered}
\operatorname{det}\left(D^{2} f\left(x_{1}, x_{2}\right)\right)= \\
\left(x_{1}^{\rho}+x_{2}^{\rho}\right)^{\frac{1}{\rho}-2} x_{1}^{2 \rho-2} x_{2}^{2 \rho-2} \operatorname{det}\left(\begin{array}{cc}
(\rho-1) & (1-\rho) \\
(1-\rho) & (\rho-1)
\end{array}\right)=0 .
\end{gathered}
$$

$D^{2} f\left(x_{1}, x_{2}\right)$ is therefore negative semidefinite if $\rho<1$ and positive semidefinite if $\rho>1$.

## Definiteness and comparative statics

Consider the unconstrained optimization of choosing $y \in \mathbb{R}$ to reach the highest possible value of:

$$
f(y ; x),
$$

where $x \in \mathbb{R}$ is an exogenous variable. Write the problem of maximizing $y: n$ as follows:

$$
\max _{y} f(y ; x)
$$

The first order condition for optimum at $(\hat{y}, \hat{x})$ is:

$$
\frac{\partial f}{\partial y}(\widehat{y} ; \widehat{x})=0
$$

A sufficient condition for local maximum is obtained from Taylor's theorem:

$$
f(\widehat{y}+d y ; \widehat{x})-f(\widehat{y} ; \widehat{x})=\frac{\partial f}{\partial y}(\widehat{y} ; \widehat{x}) d y+\frac{1}{2} \frac{\partial^{2} f}{\partial y \partial y}(\widehat{y} ; \widehat{x})(d y)^{2}+\text { h.o.t. }
$$

If

$$
\frac{\partial^{2} f}{\partial y \partial y}(\widehat{y} ; \widehat{x})<0
$$

then $f$ has a local maximum at $(\widehat{y} ; \widehat{x})$.
Note that then also the function

$$
\frac{\partial f}{\partial y}(\widehat{y} ; \widehat{x})
$$

has a non-zero derivative w.r.t. the endogenous variable at $(\widehat{y} ; \widehat{x})$ and we can apply the implicit function theorem $y$ to get the optimal $y$ as a function of $x$.

Since

$$
\frac{\partial f}{\partial y}(y(x) ; x)=0 .
$$

for all $x$ near $\widehat{x}$, we get:

$$
\frac{\partial^{2} f(\widehat{y} ; \widehat{x})}{\partial y \partial y} d y+\frac{\partial^{2} f(\widehat{y} ; \widehat{x})}{\partial y \partial x} d x=0
$$

or

$$
\frac{d y}{d x}=-\frac{\frac{\partial^{2} f(\hat{y}, \hat{x})}{\partial y \partial x}}{\frac{\partial^{2} f(\hat{y}, \hat{x})}{\partial y \partial y}} .
$$

Since $\frac{\partial^{2} f(\hat{y}, \hat{x})}{\partial y \partial y}<0$ by second-order condition for optimum, we see that $\frac{d y}{d x}$ has the same sign as $\frac{\partial^{2} f(\hat{y} ; \widehat{x})}{\partial y \partial x}$.
Example 1 (Optimal monopoly production). Let $q$ be the output by the monopolist. Denote the inverse demand by $p(q ; \alpha)$ and assume that it is twice differentiable and downward sloping $\frac{\partial p(q ; \alpha)}{\partial q}<0$ for all $q>0$. Let $\alpha$ be a demand shifting variable with such as the income of the consumers and $\frac{\partial p(q ; \alpha)}{\partial \alpha}>0$ for all $q>0$.

Denote the cost function by $c(q ; \beta)$. Assume that the cost is increasing $q$ and also that the marginal cost is increasing in $q$. Let $\beta$ be a cost shifting
exogenous variable e.g. price of oil and assume that $\frac{\partial c(q ; \beta)}{\partial \beta}<0$ for all $q>0$ The monopolist's maximization problem is then:

$$
\max _{q} \pi(q ; \alpha, \beta)=q p(q ; \alpha)-c(q ; \beta)
$$

First-order condition for optimality:

$$
D \pi(q ; \alpha, c)=p(q ; \alpha)+q \frac{\partial p(q ; \alpha)}{\partial q}-\frac{\partial c(q ; \beta)}{\partial q}=0 .
$$

Second-order condition:

$$
H \pi(q)<0
$$

If $p^{\prime}(q ; \alpha)$ is decreasing in $q$, then the second derivative is negative for all $q$.
How does the optimal output change when $\alpha$ or $c$ changes? By the previous result, the sign of the change in the endogenous variable depends on the signs of

$$
\frac{\partial^{2} \pi(q ; \alpha,)}{\partial q \partial \alpha}, \frac{\partial^{2} \pi(q ; \alpha, \beta)}{\partial q \partial \beta}
$$

Notice that we also need to know the signs of $\frac{\partial^{2} p(q ; \alpha)}{\partial q \partial \alpha}$ and $\frac{\partial^{2} c(q ; \beta)}{\partial q \partial \beta}$ to determine how optimal output changes.

