MS-A0503 First course in probability and statistics

5B More about Bayesian inference

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How to use the posterior distribution

Congratulations, you have the posterior distribution of the unknown parameter Θ . How can you use the distribution?

Like any distribution, you can use it in many ways, depending on

- 1. what question you want to answer
- 2. what is convenient to calculate.

Some typical uses:

- <u>mode</u> of the posterior distribution = where it is maximized
- mean of the posterior distribution = probability-weighted average
- <u>median</u> of the posterior distribution = 50% probability below
- credible interval, containing e.g. 95% of posterior probability
- report/visualize the <u>full</u> posterior distribution
- predictions of future data, based on posterior

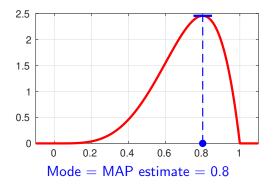
Next, we will look closer into each alternative.

Posterior mode (MAP = Maximum A Posteriori estimate)

Unknown coin, uniform prior, observed 4 heads, 1 tails. Posterior is Beta(5,2), with density (for $0 \le \theta \le 1$)

$$f(\theta \,|\, \vec{x}) = 30\theta^4(1-\theta).$$

To find the mode, inspect zeros of derivative, and ends of interval. (The normalizing constant 30 plays no role in the maximization, so we could as well use the unnormalized posterior. Also compare to ML estimate.)



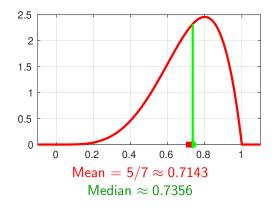
Posterior mean and median

Unknown coin, uniform prior, observed 4 heads, 1 tails. Posterior is Beta(5,2), with density (for $0 \le \theta \le 1$)

$$f(heta \,|\, ec{x}) = 30 heta^4(1- heta).$$

For mean, do the integral.

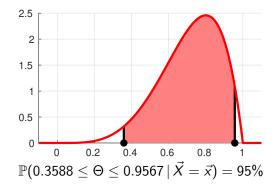
For median, solve where CDF=0.5. (R qbeta / Matlab betainv)



Credible interval

Unknown coin, uniform prior, observed 4 heads, 1 tails. Posterior is Beta(5,2), with density (for $0 \le \theta \le 1$)

$$f(\theta \,|\, \vec{x}) = 30\theta^4(1-\theta).$$



Credible interval from tables

To get the quantiles (points where CDF has desired value), in the real world you would use the computer (Matlab/Octave: betainv, R: qbeta). On this course you can also use tables.

	Ь									
а	1	2	3	4	5	6	7	8	9	10
1	0.9750	0.8419	0.7076	0.6024	0.5218	0.4593	0.4096	0.3694	0.3363	0.3085
2	0.9874	0.9057	0.8059	0.7164	0.6412	0.5787	0.5265	0.4825	0.4450	0.4128
3	0.9916	0.9324	0.8534	0.7772	0.7096	0.6509	0.6001	0.5561	0.5178	0.4841
4	0.9937	0.9473	0.8819	0.8159	0.7551	0.7007	0.6525	0.6097	0.5719	0.5381
5	0.9949	0.9567	0.9010	0.8430	0.7880	0.7376	0.6921	0.6511	0.6143	0.5810
6	0.9958	0.9633	0.9148	0.8630	0.8129	0.7662	0.7233	0.6842	0.6486	0.6162
7	0.9964	0.9681	0.9251	0.8784	0.8325	0.7891	0.7487	0.7114	0.6771	0.6457
8	0.9968	0.9719	0.9333	0.8907	0.8483	0.8078	0.7696	0.7341	0.7012	0.6708
9	0.9972	0.9748	0.9398	0.9008	0.8614	0.8234	0.7873	0.7535	0.7219	0.6924
10	0.9975	0.9772	0.9451	0.9091	0.8724	0.8366	0.8025	0.7702	0.7398	0.7114

0.975-quantiles of Beta(a, b)

Prediction of future data

- The posterior distribution of Θ is our best knowledge of what the parameter value can be (combining prior and data).
- Usually the posterior distribution is **not** a single point. This openly shows our uncertainty; we do not pretend that we know the parameter value exactly.
- But the more data we obtain, the more precise the posterior becomes.

Question. After seeing five observations $\vec{x} = (1, 1, 1, 1, 0)$, we have the posterior $\Theta \sim \text{Beta}(5, 2)$. What can we say about the <u>next</u> observation?

Answer. We form the (posterior) predictive distribution for it, applying law of total probability (= consider all possibilities and add up).

Prediction of future data (coin example)

We have five observations \vec{x} , and wish to predict next observation Y. From the law of total probability, we have

$$f_{Y|\vec{X}}(y \mid \vec{x}) = \int f(y \mid \theta) f(\theta \mid \vec{x}) d\theta.$$

Different values of θ give different predictions for Y. These predictions are <u>averaged</u>, weighted by the posterior density of Θ .

This gives our best understanding of Y, considering what we now know about Θ .

- We do **not** choose just one value of θ, perhaps the "most probable" one, and use that as the probability of Y = 1. That might give quite erroneous predictions.
- We do not even reject the 5% tails; they are included in the calculation (and for good reason: they might actually affect the prediction, see following).

Prediction of future data (coin example)

5 observations $\vec{x} = (1, 1, 1, 1, 0)$, predict next observation Y.

- stochastic model as before: $\mathbb{P}(Y = 1 | \Theta = \theta) = \theta$.
- posterior for Θ is Beta(5,2).

So calculate:

$$\mathbb{P}(Y = 1 | \vec{X} = \vec{x}) = f_{Y|\vec{X}}(1 | \vec{x})$$

$$= \int f_{Y|\Theta}(1 | \theta) f_{\Theta|\vec{X}}(\theta | \vec{x}) d\theta$$

$$= \int_0^1 \theta \cdot 30\theta^4 (1 - \theta) d\theta$$

$$= 30 \int_0^1 (\theta^5 - \theta^6) d\theta$$

$$= 30 \cdot \left(\frac{1}{6} - \frac{1}{7}\right) \approx \mathbf{0.7143}.$$

For predicting <u>one</u> more data point, our probability is simply the <u>posterior</u> mean of Θ . But don't get too carried away ...

Prediction of future data (more predictions)

5 observations $\vec{x} = (1, 1, 1, 1, 0)$, calculate probability for $\vec{Y} = (1, 1, 1)$ (that next three are heads).

- stochastic model $\mathbb{P}(\vec{Y} = (1, 1, 1) | \Theta = \theta) = \theta^3$.
- posterior for Θ is Beta(5,2).

$$\mathbb{P}(\vec{Y} = (1,1,1) | \vec{X} = \vec{x}) = f_{\vec{Y}|\vec{X}}(1,1,1|\vec{x})$$
$$= \int f_{\vec{Y}|\Theta}(1,1,1|\theta) f_{\Theta|\vec{X}}(\theta|\vec{x}) d\theta$$
$$= \int_0^1 \theta^3 \cdot 30\theta^4 (1-\theta) d\theta$$
$$= 30 \int_0^1 (\theta^7 - \theta^8) d\theta$$
$$= 30 \cdot \left(\frac{1}{8} - \frac{1}{9}\right) \approx \mathbf{0.4167}.$$

It is <u>not</u> $0.7143^3 \approx 0.3645$, but bigger. Note where the cube goes.

Prediction of future data - drastic effect of uncertainty

Being honest about our uncertainty of Θ can have a big effect on predictive distributions.

We could do this with the continuous- θ coin, but let us do a simpler **discrete** example.

Consider the following two models:

- Model A: We have a fair coin, $\Theta = 0.5$ certainly.
- Model B: We have a coin that might be unfair: Θ is either 0, 0.5 or 1, with probabilities 0.01, 0.98, 0.01 respectively.

For predicting <u>one</u> result, the models are equivalent. Each says that the next result is heads with 50% probability.

For predicting <u>next 100</u> results, the models disagree strongly.

Prediction of future data — drastic effect of uncertainty

What is the probability for the next 100 tosses to be all heads?

Model A: We know the coin is fair ($\Theta = 0.5$).

- Number of heads has Bin(100, 0.5) distribution, so
- 100 heads with probability $1/2^{100} \approx 8 \cdot 10^{-31}$

Model B: Value of Θ is 0, 0.5 or 1, with probabilities 0.01, 0.98, 0.01. (This could be our posterior from a small number of experiments.)

• By law of total probability, prob. of 100 heads is

 $(0.01 \cdot 0) + (0.98 \cdot 8 \cdot 10^{-31}) + (0.01 \cdot 1) \approx 0.01$

Observe: If Model B is the best we know, then

- Using just the mode ($\theta = 0.5$) would go wrong
- Using just the mean (heta=0.5) would go wrong
- Rejecting "5% tails" would get rid of the two extreme possibilities, and would go wrong

Keep the uncertainty in your calculations and you get more truthful results!

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Multiple categories

We worked with the <u>binary model</u>: data were 0-1-valued (or their counts), and we had a single <u>probability parameter</u>, discrete or continuous.

Next we consider sequences of categorical (nominal) data that have several categories (more than two).

Examples:

- Rolls of a loaded die (3,6,6,2,6,1,3,4,6,6)
- Party stances in a sample (A,B,A,A,C,B,A,A,C,C)
- DNA sequence with four bases chosen randomly GTCTACCAG...
- Text, as sequence of letters, according to language-specific frequency (Scrabble) t, h, e, space, q, u, i, c, k
- Text, as sequence of words, each word chosen randomly with some probabilities (the, quick, brown, fox, jumped, over, the, lazy, dog)

You can view the data either as a <u>sequence</u> of categorical variables, or as a vector of counts of the different values.

Multinomial model

- *n* independent observations (X_1, X_2, \ldots, X_n) .
- Each X_i from the same discrete distribution over k possibilies
- The distribution has k probability parameters $\vec{p} = (p_1, p_2, \dots, p_k)$
- We can treat the probabilities as unknown, a random vector $\vec{P} = (P_1, P_2, \dots, P_k)$

We can use the familiar methods:

- Assume a prior distribution $f_{\vec{P}}(\vec{p})$
- Assume a stochastic model $f(x | \vec{p})$ (likelihood)
- After observations, work out posterior $f(\vec{p} | x)$

Stochastic model — Three-category example

A large population contains supporters of three parties A, B, C with proportions $\vec{p} = (p, q, r) = (0.5, 0.3, 0.2)$.

A random sample of n = 10 people is taken. Each person sampled has the probabilities \vec{p} for the three parties.

Two questions:

- What kinds of (ordered) <u>sequences</u> are we likely to observe? example: AAAAAAAAAA or AAAABBBBCC
- What kinds of <u>count vectors</u> are we likely to observe? example: (10,0,0) or (4,4,2)

For example,

- $\mathbb{P}(AAAAAAAAA) = p^{10} \approx 0.000977$ Small
- $\mathbb{P}(AAAABBBBBCC) = p^4 q^4 r^2 \approx 0.000020$ Smaller!?

Stochastic model — Three-category example

From elementary combinatorics, we know there are $3^{10} = 59049$ different 10-person strings from three letters. Let us list them, grouped by the counts of A,B,C. Recall (p, q, r) = (0.5, 0.3, 0.2).

sequence	letter counts	$\mathbb{P}(sequence)$	
AAAAAAAAA	(10, 0, 0)	$p^{10} = 0.000977$	$}1$ sequence
AAAABBBBBCC	(4, 4, 2)	$p^4 q^4 r^2 = 0.000020$	
BBCAABBAAC	(4, 4, 2)	$p^4 q^4 r^2 = 0.000020$	
AABCCAABBB	(4, 4, 2)	$p^4 q^4 r^2 = 0.000020$	3150 seq.
CCBBBBAAAA	(4, 4, 2)	$p^4 q^4 r^2 = 0.000020$	J
СССССССССС	(0,0,10)	$r^{10} = 0.0000001$	}1 sequence

$$\begin{split} \mathbb{P}(\text{counts are } 10,0,0) &= 1 \times 0.000977 \approx 0.1\% \\ \mathbb{P}(\text{counts are } 4,4,2) &= 3150 \times 0.000020 = 0.0638 \approx 6.4\% \end{split}$$

Interlude — multinomial coefficients

Where did we get 3150 on the previous slide?

It is a <u>multinomial coefficient</u>, which tells: in how many ways can you order 4 A's, 4 B's and 2 C's, into a sequence of 10 letters.

Count the ways by the combinatorial product rule (Ross's "basic principle of counting") and the binomial coefficient:

- From the 10 places, choose 4 for the A: $\binom{10}{4} = 210$ ways
- From the remaining 6 places, choose 4 for B: $\binom{6}{4} = 15$ ways
- (From the remaining 2 places, choose 2 for C: $\binom{2}{2} = 1$ ways) Product rule: $210 \cdot 15 \cdot 1 = 3150$ ways of placing the letters.

This can be written as the multinomial coefficient

$$\binom{10}{4,4,2} = \frac{10!}{4!\ 4!\ 2!} = 3150.$$

All possible count vectors, and their probabilities

 $3^{10} = 59049$ different sequences, but 66 different count vectors.

(5,3,2)	0.0851	(2,5,3)	0.0122	(4,0,6)	0.000840
(6,2,2)	0.0709	(2,4,4)	0.0102	(2,8,0)	0.000738
(6, 3, 1)	0.0709	(4,6,0)	0.0096	(1,3,6)	0.000726
(4,4,2)	0.0638	(2,6,2)	0.0092	(1,8,1)	0.000590
(5, 4, 1)	0.0638	(3,2,5)	0.0091	(2,1,7)	0.000346
(5,2,3)	0.0567	(4,1,5)	0.0076	(0,6,4)	0.000245
(4,3,3)	0.0567	(7,0,3)	0.0075	(0,7,3)	0.000210
(7, 2, 1)	0.0506	(8,0,2)	0.0070	(1,2,7)	0.000207
(4, 5, 1)	0.0383	(9,1,0)	0.0059	(0,5,5)	0.000196
(3,4,3)	0.0340	(2,3,5)	0.0054	(3,0,7)	0.000192
(7, 1, 2)	0.0338	(6,0,4)	0.0053	(0,8,2)	0.000118
(6,1,3)	0.0315	(2,7,1)	0.0039	(0,4,6)	0.000109
(3,5,2)	0.0306	(9,0,1)	0.0039	(1,9,0)	0.000098
(4,2,4)	0.0284	(3,7,0)	0.0033	(0,3,7)	0.000041
(6,4,0)	0.0266	(5,0,5)	0.0025	(0,9,1)	0.000039
(7,3,0)	0.0253	(1,6,3)	0.0024	(1,1,8)	0.000035
(3,3,4)	0.0227	(1,5,4)	0.0024	(2,0,8)	0.000029
(8, 1, 1)	0.0211	(3,1,6)	0.0020	(0,2,8)	0.000010
(5,5,0)	0.0191	(2,2,6)	0.0018	(0,10,0)	0.000006
(5,1,4)	0.0189	(1,4,5)	0.0016	(1,0,9)	0.000003
(8,2,0)	0.0158	(1,7,2)	0.0016	(0,1,9)	0.000002
(3,6,1)	0.0153	(10,0,0)	0.000977	(0,0,10)	0.000000

Multinomial model — A discrete prior

For the probability parameter vector \vec{P} , in different situations we can have different kinds of priors.

Sometimes we just have <u>a few</u> possible values of the vector, perhaps just two, so the prior distribution is discrete.

E.g. we have just two kinds of dice in a bag: 9 fair and 1 loaded, and we know the loading. A randomly chosen die is

- with prob. 0.9 fair, with $\vec{p} = (\frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6}, \frac{1}{6})$
- with prob. 0.1 loaded, with $\vec{p} = (0.1, 0.1, 0.1, 0.1, 0.1, 0.5)$

Pick a random die, roll it four times with results (3, 2, 6, 6). The likelihoods for the two possible parameter values are

- for a fair die: $(\frac{1}{6})^4 \approx 0.00077$
- for the loaded die: $0.1 \cdot 0.1 \cdot 0.5 \cdot 0.5 = 0.00250$

After this observation, we would have increased posterior probability for the die to be the loaded one. (But not certainty!)

Multinomial model — A continuous prior

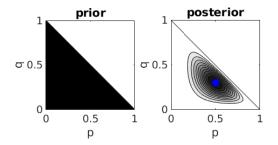
Or perhaps the values of the three probability parameters (p, q, r) are unknown real numbers. How do we handle this situation?

- Certainly all three are in the interval [0, 1].
- They are not three freely chosen parameters, because we must have p + q + r = 1.
- We can consider a two-element parameter vector (p, q), and then r = 1 (p + q).
- We need p ≥ 0 and q ≥ 0 and p + q ≤ 1. So (p, q) is constrained to be in a triangular area. (Picture!)
- Let the prior be the uniform density over the triangle,

$$f_{P,Q}(p,q)=2 \qquad \quad \text{if $p,q\geq 0$ and $p+q\leq 1$.}$$

• We now have the likelihood and the prior, so we can proceed with Bayesian inference.

Multinomial model — Inference



After observing counts (5, 3, 2), the posterior density of (P, Q) is

$$f(p,q \,|\, \vec{x}) = c \cdot p^5 q^3 (1-p-q)^2$$

in the triangle, and c is again normalizing constant. We can use the posterior density to compute posterior mode, posterior mean, 95% credible <u>region</u>, predictions etc. Posterior mode here shown as blue dot. Credible area: live demo

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Some benefits of the Bayesian approach

If you are willing to treat Θ as a random variable (and assign a prior distribution to it), you gain:

- Mathematical unification. "Parameters" and "observations" are unified as "quantities" that follow the same mathematical laws of probability.
- **General applicability.** With the same framework, you can calculate posteriors
 - for small data, even *n* = 1, where e.g. "normal approximations" would not apply at all
 - for big data
 - for non-normal data, e.g. exponential observations
 - for more complicated models
- Full posterior distribution of Θ. It gives you a richer understanding of Θ's possible values than just a single point estimate or interval estimate. You can inspect it visually, and ask and answer any questions like mean, mode, median, probability of this interval ...

Choice of prior

Sometimes people are worried about the apparent subjectivity of Bayesian inference. If you <u>want</u> to report a certain posterior distribution you like, you could <u>choose</u> your prior so that you get the posterior you wanted?

- You should be honest in making your prior to be a fairly good representation of <u>what is known</u> about Θ before the data.
- Uniform priors often work out nice. Not always, in complicated models.
- Beware of assigning <u>zero</u> density to some parameter values that might actually be true. Zero prior leads to zero posterior, whatever your data are.
- With lots of data, the effect of the prior <u>diminishes</u> as the "data speaks for itself".
- When reporting your results, <u>report</u> the model and prior you used. Then your results are completely objective: anyone using <u>that</u> prior will get the <u>same</u> posterior.

Next week: Significance tests...