ELEC-E8107 Stochastics models, estimation and Control (5 p)

Applications of the **state** estimation: **Kalman** filtering

- Tracking/ surveillance
- Control systems, power systems, failure detection
- Navigation and trajectory determination, robotics
 Probabilistic robotics
- Signal processing, image processing, communication, operations research, econometric systems
- Remote sensing, geophysical research, Biomedical systems

Duality of estimation and control, stochastic optimal control



Course arrangements

- Lectures: <u>Arto.Visala@aalto.fi</u> on Tuesday 14:15-16:00
- Exercises and home assignments: <u>Issouf.ouattara</u> <u>@aalto.fi</u> on Tuesdays 8:15-10:00
- Slides mainly from textbook of Bar-Shalom et al
- In autumn 2022, evaluation consists of Exam (100%)
- In exam you can get extra points corresponding one grade by submitting in time valid Home assignments given in connection of exercises.
- It is possible to get the excellent grade without Home assignments, but it is not easy.



Field robotics; autonomous vehicles Data from ATV platforms used is in examples







ELEC-E8107 Stochastics models, estimation and control Lecture 1: Statistics and stochastics

- Gradient, Jacobian and Hessian; Eigenvalues, Eigenvectors, and Quadratic Forms
- Gaussian Random Variables; pdf, Mean and Covariance; Joint and Conditional Random Variables
- Fundamental Equations of Linear Estimation
- Stochastic processes; Correlation; White noise process
- Random Sequences, Markov Processes; Markov Sequences and Markov Chains

Statistics and stochastics in Kalman



Gradient, Jacobian, Hessian





Hessian

• Hessian of the scalar function

$$\phi_{xx}(x) \stackrel{\Delta}{=} \frac{\partial^2 \phi(x)}{\partial x^2} = \nabla_x \nabla'_x \phi(x) = \begin{bmatrix} \frac{\partial^2 \phi}{\partial x_1 \partial x_1} & \cdots & \frac{\partial^2 \phi}{\partial x_1 \partial x_n} \\ \vdots & \ddots & \vdots \\ \frac{\partial^2 \phi}{\partial x_n \partial x_1} & \cdots & \frac{\partial^2 \phi}{\partial x_n \partial x_n} \end{bmatrix}$$



Eigenvalues, Eigenvectors, and Quadratic Forms

A nxn matrix, eigenvectors and -values

$$\begin{array}{ll} Au_i = \lambda_i u_i & i = 1, \ldots, n \\ \text{Determinant} & |A| = \prod_{i=1}^n \lambda_i \\ \text{Trace} & \operatorname{tr}(A) = \sum_{i=1}^n \lambda_i \end{array}$$

Quadratic form
$$q = x'Ax$$



A positive (semi)definite quadratic form

 $A > 0 \qquad \iff \qquad x'Ax > 0 \quad \forall x \neq 0$ $A \ge 0 \qquad \iff \qquad x'Ax \ge 0 \quad \forall x \neq 0$

The inequality of two matrices is defined as follows: the matrix A is smaller (not larger) than the matrix B if and only if the difference B - A is positive (semi) definite



Gaussian pdf, random variable, mean and covariance

 Scalar Gaussian or Normal pdf

$$p(x) = \mathcal{N}(x; \bar{x}, \sigma^2) \stackrel{\Delta}{=} \frac{1}{\sqrt{2\pi\sigma}} e^{-\frac{(x-\bar{x})^2}{2\sigma^2}}$$

$$x \sim \mathcal{N}(\bar{x}, \sigma^2)$$

$$\mathcal{N}(x;\bar{x},P) \stackrel{\Delta}{=} |2\pi P|^{-1/2} e^{-\frac{1}{2}(x-\bar{x})'P^{-1}(x-\bar{x})}$$

- Mean or expected value
- Covariance P

$$\bar{x} = E[x]$$

$$P = E[(x - \bar{x})(x - \bar{x})']$$



Mean, average, first moment

$$E[x] = \int_{-\infty}^{\infty} x p(x) \, dx \stackrel{\Delta}{=} \bar{x}$$

The nth moment is

$$E[x^n] = \int_{-\infty}^{\infty} x^n p(x) \, dx$$

The second central moment or variance

$$\operatorname{var}(x) \stackrel{\Delta}{=} E[(x - \bar{x})^2] =$$

$$\int_{-\infty}^{\infty} (x - \bar{x})^2 p(x) \, dx = E[x^2] - (\bar{x})^2 \stackrel{\Delta}{=} \sigma_x^2$$

Mean Square MS

$$E[x^2] = [E(x)]^2 + \operatorname{var}(x) = \bar{x}^2 + \sigma_x^2$$



The Central Limit Theorem

If the sequence x_i , i = 1, ..., consists of independent random variables, then under some reasonably mild conditions the pdf of the sum

 $z_n = \frac{1}{\sqrt{n}} \sum_{i=1}^n x_i$ will tend to a Gaussian pdf as $n \to \infty$.

The Law of Large Numbers

states loosely that the sum of a large number of random variables tends, under some fairly nonrestrictive conditions, to its expected value.



Joint and Conditional Gaussian random variables, Conditional pdf of x given z.

$$y = \begin{bmatrix} x \\ z \end{bmatrix} \qquad p(x,z) = p(y) = \mathcal{N}(y;\bar{y}, P_{yy}) \qquad \bar{y} = \begin{bmatrix} x \\ \bar{z} \end{bmatrix}$$
$$P_{xx} = \operatorname{cov}(x) = E[(x - \bar{x})(x - \bar{x})'] \qquad P_{zz} = \operatorname{cov}(z) = E[(z - \bar{z})(z - \bar{z})'] \qquad P_{yy} = \begin{bmatrix} P_{xx} & P_{xz} \\ P_{zx} & P_{zz} \end{bmatrix}$$
$$P_{xz} = \operatorname{cov}(x,z) = E[(x - \bar{x})(z - \bar{z})'] = P'_{zx} \qquad P_{yy} = \begin{bmatrix} P_{xx} & P_{xz} \\ P_{zx} & P_{zz} \end{bmatrix}$$
$$p(x|z) = \frac{p(x,z)}{p(z)} = \frac{|2\pi P_{yy}|^{-1/2} e^{-\frac{1}{2}(y - \bar{y})' P_{yy}^{-1}(y - \bar{y})}}{|2\pi P_{zz}|^{-1/2} e^{-\frac{1}{2}(z - \bar{z})' P_{zz}^{-1}(z - \bar{z})}}$$

New zero mean random variables, the exponent q of Gaussian conditional pdf becomes

$$\xi \stackrel{\Delta}{=} x - \bar{x} \qquad \qquad q = \begin{bmatrix} \xi \\ \zeta \end{bmatrix} \begin{bmatrix} P_{xx} & P_{xz} \\ P_{zx} & P_{zz} \end{bmatrix} \begin{bmatrix} \xi \\ \zeta \end{bmatrix} - \zeta' P_{zz}^{-1} \zeta$$
$$\zeta \stackrel{\Delta}{=} z - \bar{z} \qquad \qquad = \begin{bmatrix} \xi \\ \zeta \end{bmatrix}' \begin{bmatrix} T_{xx} & T_{xz} \\ T_{zx} & T_{zz} \end{bmatrix} \begin{bmatrix} \xi \\ \zeta \end{bmatrix} - \zeta' P_{zz}^{-1} \zeta$$



continues, matrix inversion lemma is utilized; trick $TT^{-1} = I$ $T_{xx}^{-1} = P_{xx} - P_{xz}P_{zz}^{-1}P_{zx}$ $P_{zz}^{-1} = T_{zz} - T_{zx}T_{xx}^{-1}T_{xz}$ $T_{xx}^{-1}T_{xz} = -P_{xz}P_{zz}^{-1}$

$$q = \xi' T_{xx} \xi + \xi' T_{xz} \zeta + \zeta' T_{zx} \xi + \zeta' T_{zz} \zeta - \zeta' P_{zz}^{-1} \zeta$$

$$= (\xi + T_{xx}^{-1} T_{xz} \zeta)' T_{xx} (\xi + T_{xx}^{-1} T_{xz} \zeta) + \zeta' (T_{zz} - T_{zx} T_{xx}^{-1} T_{xz}) \zeta - \zeta' P_{zz}^{-1} \zeta$$

$$= (\xi + T_{xx}^{-1} T_{xz} \zeta)' T_{xx} (\xi + T_{xx}^{-1} T_{xz} \zeta)$$
(1.4.14-15)

completion of the squares in exponent of $\xi + T_{xx}^{-1}T_{xz}\zeta = x - \bar{x} - P_{xz}P_{zz}^{-1}(z - \bar{z})$ the conditional Gaussian pdf

$$E(x|z) \triangleq \hat{x} = \bar{x} + P_{xz}P_{zz}^{-1}(z-\bar{z})$$
$$cov(x|z) \triangleq P_{xx|z} = T_{xx}^{-1} = P_{xx} - P_{xz}P_{zz}^{-1}P_{zx}$$

Fundamental equations of linear estimation



Estimation of Gaussian random vectors

x and z jointly Gaussian z is the measurement x random variable to be estimated $y \triangleq \begin{bmatrix} x \\ z \end{bmatrix}$ $\bar{y} = \begin{bmatrix} \bar{x} \\ \bar{z} \end{bmatrix}$ $P_{yy} = \begin{bmatrix} P_{xx} & P_{xz} \\ P_{zx} & P_{zz} \end{bmatrix}$

 $y \sim \mathcal{N}[\bar{y}, P_{yy}]$ $P_{xx} = E[(x - \bar{x})(x - \bar{x})']$ $P_{xz} = E[(x - \bar{x})(z - \bar{z})'] = P'_{zx}$

Next slide set: The **MMSE Minimum Mean Square Error** –estimator is the conditional mean of x given z, for linear Gaussian case, also **Maximum a Posteriori MAP** -estimator

$$\hat{x} \stackrel{\Delta}{=} E[x|z] = \bar{x} + P_{xz}P_{zz}^{-1}(z - \bar{z})$$

$$P_{xx|z} \stackrel{\Delta}{=} E[(x - \hat{x})(x - \hat{x})'|z] = P_{xx} - P_{xz}P_{zz}^{-1}P_{zx}$$



Fundamental equations of linear estimation - Interpretations

- A priori estimate is updated/corrected on the basis of measurement information in calculation of a posteriori estimate
- Correction gain depends directly on P_{xz} the crosscovariance between x and measurement z. P_{xz} must be unzero in order measurements contain information in general about the state x.
- Correction effect depends inversely proportional on P_{zz} . The better measurements, the 'smaller' covariance, the bigger the correction gain.



Random process, stochastic process

A scalar random variable is a (real) number x determined by the outcome ω of a random experiment

$$x = x(\omega)$$

A (scalar) *random process or a stochastic process* is a function of time determined by the outcome of a random experiment

$$x(t) = x(t,\omega)$$

This is a family or ensemble of functions of time, in general different for each outcome $\boldsymbol{\omega}$

mean or ensemble average

$$\bar{x}(t) = E[x(t)] = \int_{-\infty}^{\infty} \xi \ p_{x(t)}(\xi) \, d\xi$$



its autocorrelation

$$R(t_1, t_2) \stackrel{\Delta}{=} E[x(t_1)x(t_2)] =$$
$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \xi \eta \ p_{x(t_1), x(t_2)}(\xi, \eta) \ d\xi d\eta$$

autocovariance of this random process

$$V(t_1, t_2) \stackrel{\Delta}{=} E[[x(t_1) - \bar{x}(t_1)][x(t_2) - \bar{x}(t_2)]]$$

= $R(t_1, t_2) - \bar{x}(t_1)\bar{x}(t_2)$

Stationarity process

$$R(t_1, t_2) = R(t_1 - t_2) = R(\tau)$$

The *power spectrum* or *power spectral density* of a *stationary* random process is the Fourier transform

$$S(\omega) = \mathcal{F}\{R(\tau)\} = \int_{-\infty}^{\infty} e^{-j\omega\tau} R(\tau) d\tau$$



White Noise

A (not necessarily stationary) random process whose autocovariance is zero for any two different times is called *white noise*

$$V(t_1, t_2) = \sigma^2(t_1)\delta(t_1 - t_2) = S_0(t_1)\delta(t_1 - t_2)$$

 $\delta(\cdot)$ the **Dirac (impulse) delta function**

A stationary zero-mean white process

$$R(\tau) = E[x(t+\tau)x(t)] = S_0\delta(\tau)$$

A nonstationary zero-mean white process x(t)

$$E[x(t_1)x(t_2)] = S_0(t_1)\delta(t_1 - t_2)$$



Random Walk and the Wiener Process

The *Wiener random process* (or Wiener-Levy or Brownian motion) is a limiting form of the *random walk*: the sum of independent steps of size $s \rightarrow 0$, equiprobable in each direction, taken at intervals $\Delta \rightarrow 0$ such that

$$\frac{\delta}{\sqrt{\Delta}} \to \sqrt{\alpha}$$

This yields a stochastic process $\mathbf{w}(t)$ with the following pdf [assuming $\mathbf{w}(0) = 0$],

$$p[\mathbf{w}(t)] = \mathcal{N}[\mathbf{w}(t); 0, \alpha t]$$

the *Wiener process* is *nonstationary*. It relates to the zero-mean white noise, denoted here as n(t), as follows

$$\mathbf{w}(t) = \int_0^t n(\tau) d\tau \qquad E[n(t_1)n(t_2)] = \alpha \delta(t_1 - t_2)$$
$$d\mathbf{w}(t) = n(t) dt$$



Stochastic sequences, Markov property

Markov processes are defined by Markov property

 $p[x(t)|x(\tau), \tau \le t_1] = p[x(t)|x(t_1)] \qquad \forall t > t_1$

the past up to any t_1 is *fully characterized* by the value of the process at t_1

"The future is independent of the past if the present is known."

The Wiener process is Markov

Furthermore, the state x(t) of a (possibly time-varying) dynamic system driven by white noise n(t), is a Markov process

$$\dot{x}(t) = f[t, x(t), n(t)]$$



Random Sequences and Markov Sequences

Random Sequences,

$$X^k = \{x(j)\}_{j=1}^k \qquad k = 1, 2, \dots$$

a random sequence is Markov Sequences if

 $p[x(k)|X^j] = p[x(k)|x(j)] \qquad \forall k > j$

The (real-valued) zero-mean sequence v(j), j = 1, ..., is a discrete-time white noise (a white sequence) if

 $E[v(k)v(j)] = q(k)\delta_{kj}$

where the *Kronecker delta function* q(k) denotes its variance, q(k) = q stationaty

$$\delta_{kj} = \begin{cases} 1 & \text{if } k = j \\ 0 & \text{if } k \neq j \end{cases}$$



The state x(k) of a dynamic system excited by white noise v(k)

x(k+1) = f[k, x(k), v(k)]

is a *discrete-time Markov process* or *Markov sequence*. In general, both x(k) and v(k) are vector-valued.

The state of a linear dynamic system excited by white Gaussian noise

x(k+1) = F(k)x(k) + v(k)

is a Gauss-Markov sequence

A special case, for a scalar x,

$$x(k+1) = x(k) + v(k)$$

x becomes the *integral (sum)* of the white noise sequence terms, and is called a *discrete-time Wiener process*



Markov Chains

Define

A *Markov chain* is a special case of a Markov sequence, in which the **state space** is **discrete** and **finite**

$$x(k) \in \{x_i, i = 1, \dots, n\}$$

its characterization is given in full by the transition (jump) probabilities

$$P\{x(k) = x_j \mid x(k-1) = x_i\} \stackrel{\Delta}{=} \pi_{ij} \qquad i, j = 1, \dots, n$$

the vector
$$\begin{array}{c} \mu(k) \stackrel{\Delta}{=} [\mu_1(k), \dots, \mu_n(k)]' \\ \mu_i(k) \stackrel{\Delta}{=} P\{x(k) = x_i\} \end{array}$$

where the components are the probabilities of the chain being in state *i*. The evolution in time of (1.4.22-11) is then given by

$$\mu_i(k+1) = \sum_{j=1}^n \pi_{ji} \mu_j(k) \qquad i = 1, \dots, n$$

$$\mu(k+1) = \Pi' \mu(k) \qquad \text{with matrix notation} \qquad \Pi = [\pi_{ij}]$$

$$\text{the transition matrix of the Markov chain}$$



Markov property continues

discrete-time white noise (a white sequence) scalar E

$$E[v(k)v(j)] = q(k)\delta_{kj}$$
(1.4.22-3)

Kronecker delta function

$$\delta_{kj} = \begin{cases} 1 & \text{if } k = j \\ 0 & \text{if } k \neq j \end{cases}$$
(1.4.22-4)

The state of linear dynamic equation with white noise process

$$E[v(k)] = 0 (4.3.1-6)$$

$$E[v(k)v(j)'] = Q(k)\delta_{kj}$$
(4.3.1-7)

$$x(k+1) = F(k)x(k) + G(k)u(k) + v(k)$$
(4.3.1-9)

has the following analytical solution, next slide



The State as a Markov Process
$$x(k) = \left[\prod_{j=0}^{k-l-1} F(k-1-j)\right] x(l) + \sum_{i=l}^{k-1} \left[\prod_{j=0}^{k-i-2} F(k-1-j)\right] [G(i)u(i) + v(i)]$$
(4.3.3-1)

Thus, since v(i), i = 1, ..., k - 1, are independent of

$$X^{l} \stackrel{\Delta}{=} \{x(j)\}_{j=0}^{l}$$
(4.3.3-2)

which depend only on v(i), i = 0, ..., I - 1, one has

$$p[x(k)|X^{l}, U^{k-1}] = p[x(k)|x(l), U_{l}^{k-1}] \qquad \forall k > l$$
(4.3.3-3)

$$U_l^{k-1} \stackrel{\Delta}{=} \{u(j)\}_{j=l}^{k-1}$$
(4.3.3-4)

Thus, **the state vector is** a Markov process, or, more correctly, a **Markov sequence**.

State of a stochastic system described by a Markov process — summarizes probabilistically its past.

