GIS-E3010 Least-Squares Methods in Geoscience Lecture 2/2018

- Observation equation model
- Derivation of general solution
- Examples

Why redundant observations

Uncertainties in measurements

- Instruments
- Circumstances
- Observer
- Methods
- Difference between a mathematical model and reality
- The purpose of measurements
- Economical reasons



Redundant observations

If we have more observation than necessary, we need adjustment

Questions in adjustment calculus and design of measurements

- Repeated observations do not give the same answer or redundant observations are not consistent
- Can we detect blunders, gross errors, outliers from our data
- Can we detect systematic errors
- What is the best and the most reliable way to take into account all observations and to get the best final results
- How can we prevent the corruption of the results due to the nondetected outliers



Examples

- From height differences to heights
- From angles to angles (or shape of the triangle)
- From angles, distances and GPS-vectors to the vector between two points







Examples 2



- From angles, distances, height differences and GPS-vectors to 3D- coordinates
- From angles to 3D-coordinates



Still one example







- From angles and distances to 3D coordinates (red) and from 3D coordinates to the reference point and axis directions of the telescope
- From GPS phase observations to 3D coordinates (blue and green) and from 3D coordinates to reference points and axis direction of the telescope

Models

 Observation equation model (Gauss-Markov)

 $f_i(x_1, x_2, ..., x_u, \ell_i) = 0$ Ax - y = v

 Condition equation model

 $f_i(\ell_1, \ell_2, \dots, \ell_n) = 0$ Bv - y = 0

• General or mixed model (Gauss-Helmert)

$$f_i(x_1, x_2, ..., x_u, \ell_1, \ell_2, ..., \ell_n) = 0$$

$$A(x - x_0) + Bv - y = 0$$

Notation

- *x* is unknown parameters
- *u* is number of unknown parameter
- *n* is number of observations
- *A* is design matrix (coefficients of unknown parameters)
- y is y-vector, opposite number of calculated minus
 observed (in linear model observations and possible
 constants, when approximate values of parameters are
 zeros)
- ℓ is observation
- *v* is residual vector, adjusted minus observed
- *f* is functional model, the relation between observations and unknown parameters
- P is weight matrix

Observation equation model, linear model

$$\begin{cases} a_{10} + a_{11} \cdot x_1 + a_{12} \cdot x_2 + a_{13} \cdot x_3 + \dots + a_{1u} \cdot x_u - \ell_1 = 0 \\ a_{20} + a_{21} \cdot x_1 + a_{22} \cdot x_2 + a_{23} \cdot x_3 + \dots + a_{2u} \cdot x_u - \ell_2 = 0 \\ a_{30} + a_{31} \cdot x_1 + a_{32} \cdot x_2 + a_{33} \cdot x_3 + \dots + a_{3u} \cdot x_u - \ell_3 = 0 \\ \vdots \\ a_{n0} + a_{n1} \cdot x_1 + a_{n2} \cdot x_2 + a_{n3} \cdot x_3 + \dots + a_{nu} \cdot x_u - \ell_n = 0 \end{cases}$$

$$\begin{pmatrix} a_{10} \\ a_{20} \\ a_{30} \\ \vdots \\ a_{n0} \end{pmatrix} + \begin{pmatrix} a_{11} & a_{12} & a_{13} & \cdots & a_{1u} \\ a_{21} & a_{22} & a_{23} & \cdots & a_{2u} \\ a_{31} & a_{32} & a_{33} & \cdots & a_{3u} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ a_{n1} & a_{n2} & a_{n3} & \cdots & a_{nu} \end{pmatrix} \cdot \begin{pmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_u \end{pmatrix} - \begin{pmatrix} \ell_1 \\ \ell_2 \\ \ell_3 \\ \vdots \\ \ell_n \end{pmatrix} = \begin{pmatrix} 0 \\ 0 \\ 0 \\ \vdots \\ 0 \end{pmatrix}$$

$$\begin{cases} a_{11} \cdot x_1 + a_{12} \cdot x_2 + a_{13} \cdot x_3 + \dots + a_{1u} \cdot x_{u1} - \ell_{1_{obs}} + a_{10} = v_1 \\ a_{21} \cdot x_1 + a_{22} \cdot x_2 + a_{23} \cdot x_3 + \dots + a_{2u} \cdot x_{u2} - \ell_{2_{obs}} + a_{20} = v_2 \\ a_{31} \cdot x_1 + a_{32} \cdot x_2 + a_{33} \cdot x_3 + \dots + a_{3u} \cdot x_{u3} - \ell_{3_{obs}} + a_{30} = v_3 \\ \vdots \\ a_{n1} \cdot x_1 + a_{n2} \cdot x_2 + a_{n3} \cdot x_3 + \dots + a_{nu} \cdot x_u - \ell_{n_{obs}} + a_{n0} = v_n \end{cases}$$

$$\begin{pmatrix} a_{11} & a_{12} & a_{13} & \cdots & a_{1u} \\ a_{21} & a_{22} & a_{23} & \cdots & a_{2u} \\ a_{31} & a_{32} & a_{33} & \cdots & a_{3u} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ a_{n1} & a_{n2} & a_{n3} & \cdots & a_{nu} \end{pmatrix} \cdot \begin{pmatrix} x_1 \\ x_2 \\ x_3 \\ \vdots \\ x_u \end{pmatrix} - \begin{pmatrix} y_1 \\ y_2 \\ y_3 \\ \vdots \\ y_n \end{pmatrix} = = \begin{pmatrix} v_1 \\ v_2 \\ v_3 \\ \vdots \\ v_n \end{pmatrix}$$

Ax - y = v

Linearization with Taylor



$$f(x) = f(x_0) + \frac{d}{dx}f(x_0)\Delta x + \frac{1}{2!}\frac{d^2}{dx^2}f(x_0)\Delta x^2 + \dots + \frac{1}{(q-1)!}\frac{d^{(q-1)}}{dx^{(q-1)}}f(x_0)\Delta x^{(q-1)} + R_q(\theta,\Delta x)$$

$$R_q(\theta, \Delta x) = \frac{1}{(q)!} \frac{d^{(q)}}{dx^{(q)}} f(\theta) \Delta x^q$$

Least Squares Method in Geoscience

Linearization

$$F(x, \ell) = F(x_0, \ell_0) + \frac{\partial F}{\partial x}(x - x_0) + \frac{\partial F}{\partial \ell}(\ell - \ell_0) = 0$$

Approximate value + correction
$$-y + A(x - x_0) + Bv = 0$$

$$A = \frac{\partial F(x_0, \ell_0)}{\partial x} = \begin{pmatrix} \frac{\partial f_1}{\partial x_1} & \frac{\partial f_1}{\partial x_2} & \cdots & \frac{\partial f_1}{\partial x_u} \\ \frac{\partial f_2}{\partial x_1} & \frac{\partial f_2}{\partial x_2} & \cdots & \frac{\partial f_2}{\partial x_u} \\ \vdots & \vdots & \vdots \\ \frac{\partial f_n}{\partial x_1} & \frac{\partial f_n}{\partial x_2} & \cdots & \frac{\partial f_n}{\partial x_u} \end{pmatrix} \qquad B = \frac{\partial F(x_0, \ell_0)}{\partial \ell} = \begin{pmatrix} \frac{\partial f_1}{\partial \ell_1} & \frac{\partial f_1}{\partial \ell_2} & \cdots & \frac{\partial f_1}{\partial \ell_n} \\ \frac{\partial f_2}{\partial \ell_1} & \frac{\partial f_2}{\partial \ell_2} & \cdots & \frac{\partial f_2}{\partial \ell_n} \\ \vdots & \vdots & \vdots \\ \frac{\partial f_n}{\partial \ell_1} & \frac{\partial f_n}{\partial \ell_2} & \cdots & \frac{\partial f_n}{\partial \ell_n} \end{pmatrix}$$

Observation equation model, nonlinear model differnce of y-vector and observation ℓ

$$\begin{split} f_i(x_1, x_2, \dots, x_u, \ell_i) &\approx f_i(x_{l_0}, x_{2_0}, \dots, x_{u_0}, \ell_{i_0}) \\ &+ \frac{\partial f_i(x_{l_0}, x_{2_0}, \dots, x_{u_0}, \ell_{i_0})}{\partial x_1} \cdot (\hat{x}_1 - x_{l_0}) + \frac{\partial f_i(x_{l_0}, x_{2_0}, \dots, x_{u_0}, \ell_{i_0})}{\partial x_2} \cdot (\hat{x}_2 - x_{2_0}) + \dots + \frac{\partial f_i(x_{l_0}, x_{2_0}, \dots, x_{u_0}, \ell_{i_0})}{\partial x_u} \cdot (\hat{x}_u - x_{u_0}) \\ &+ \frac{\partial f_i(x_{l_0}, x_{2_0}, \dots, x_{u_0}, \ell_{i_0})}{\partial \ell_i} \cdot (\hat{\ell}_i - \ell_{i_0}) \\ &= 0 \end{split}$$

If ℓ_i can be directly expressed with parameters x_i , the equation above is $f_i = g(x_1, x_2, \dots, x_u) - \ell_i = 0$

Thus the last partial derivative is -1.

By substituting

$$\hat{\ell}_i - \ell_{i_{obs}} = v_i$$
 and $-f_i(x_{1_0}, x_{2_0}, \dots, x_{u_0}, \ell_{i_0}) = y_i$

We obtain

$$A(x-x_0)-y=v$$

Linearized model

$$\begin{split} f_{i}(x_{1}, x_{2}, \dots, x_{u}, \ell_{1}) &\approx f_{1}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{1_{0}}) \\ &+ \frac{\partial f_{1}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{1_{0}})}{\partial x_{1}} \cdot (\hat{x}_{1} - x_{1_{0}}) + \frac{\partial f_{1}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{1_{0}})}{\partial x_{2}} \cdot (\hat{x}_{2} - x_{2_{0}}) + \dots + \frac{\partial f_{1}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{1_{0}})}{\partial x_{u}} \cdot (\hat{x}_{u} - x_{u_{0}}) \\ &+ \frac{\partial f_{1}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{1_{0}})}{\partial \ell_{1}} \cdot (\hat{\ell}_{1} - \ell_{1_{0}}) = 0 \\ f_{2}(x_{1}, x_{2}, \dots, x_{u}, \ell_{i}) &\approx f_{2}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{2_{0}}) \\ &+ \frac{\partial f_{2}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{2_{0}})}{\partial x_{1}} \cdot (\hat{x}_{1} - x_{1_{0}}) + \frac{\partial f_{2}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{2_{0}})}{\partial x_{2}} \cdot (\hat{x}_{2} - x_{2_{0}}) + \dots + \frac{\partial f_{2}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{2_{0}})}{\partial x_{u}} \cdot (\hat{x}_{u} - x_{u_{0}}) \\ &+ \frac{\partial f_{2}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{2_{0}})}{\partial \ell_{2}} \cdot (\hat{\ell}_{2} - \ell_{2_{0}}) = 0 \\ &\vdots \\ f_{n}(x_{1}, x_{2}, \dots, x_{u}, \ell_{i}) &\approx f_{n}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{u_{0}}) \\ &+ \frac{\partial f_{n}(x_{1_{u}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{u_{0}})}{\partial \ell_{2}} \cdot (\hat{\ell}_{1} - x_{1_{0}}) + \frac{\partial f_{i}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{u_{0}})}{\partial x_{2}} \cdot (\hat{x}_{2} - x_{2_{0}}) + \dots + \frac{\partial f_{n}(x_{1_{0}}, x_{2_{0}}, \dots, x_{u_{0}}, \ell_{u_{0}})}{\partial x_{u}} \cdot (\hat{x}_{u} - x_{u_{0}}) \\ &= \frac{f_{n}(x_{1}, x_{2}, \dots, x_{u}, \ell_{i}) \approx f_{n}(x_{1}, x_{2}, \dots, x_{u_{0}}, \ell_{u_{0}})}{\partial \ell_{2}} \cdot (\hat{\ell}_{2} - \ell_{2_{0}}) = 0 \\ &= \frac{1}{3} + \frac{\partial f_{n}(x_{1}, x_{2}, \dots, x_{u_{0}}, \ell_{u_{0}})}{\partial x_{u}} \cdot (\hat{\ell}_{n} - \ell_{u_{0}}) = 0 \quad \text{Least Squares Method in Geoscience} \quad 13 \\ \end{bmatrix}$$

General solution: deterministic derivation

$$Ax - y = v$$

$$v^{T}Pv = \min$$

$$\Rightarrow (Ax - y)^{T}P(Ax - y) = \min$$

$$\Rightarrow (x^{T}A^{T}P - y^{T}P)(Ax - y) = \min$$

$$\Rightarrow x^{T}A^{T}PAx - x^{T}A^{T}Py - y^{T}PAx + y^{T}Py = \min$$

$$\Rightarrow 2x^{T}A^{T}PA - y^{T}PA - y^{T}PA = 0$$

$$\Rightarrow 2x^{T}A^{T}PA - 2y^{T}PA = 0$$

$$\Rightarrow x^{T}A^{T}PA = y^{T}PA$$

$$\Rightarrow A^{T}PAx = A^{T}Py$$
Normaaliyhtälöt
Normal equations

If we have linear form $u = x^T A y$, then $\frac{\partial u}{\partial x} = y^T A^T$ and $\frac{\partial u}{\partial y} = x^T A$ For quadratic form $q = x^T A x$, $\frac{\partial q}{\partial x} = 2x^T A$

Solution of normal equations

 $x = (A^T P A)^{-1} A^T P y$

Weighting of observations

- Weight matrix P is inverse of the covariance matrix of the observations
- Variance factor ${\sigma_0}^2$ is the variance of an observation which has the weight 1

$$P = \sigma_0^2 \Sigma^{-1}$$

Variance factor can be chosen

The solution does not depend on the choise of the variance factor σ_0^2

$$x = (A^{T} \sigma_{0}^{2} \Sigma^{-1} A)^{-1} A^{T} \sigma_{0}^{2} \Sigma^{-1} y$$
$$= \frac{1}{\sigma_{0}^{2}} (A^{T} \Sigma^{-1} A)^{-1} A^{T} \sigma_{0}^{2} \Sigma^{-1} y$$

Exercise: arithmetic mean

- What is number of equations in observation equation model?
- What is number of unknown parameters?
- Functional model?
- A-matrix?
- y-vector?
- Normal equations?
- LSQ solution?

Exercise: Linear regression



- How many observations?
 - How many unknown parameters?
 - Functional model ?
 - A-matrix?
 - y-vector?

Exercise: levelling network



- Height differences between points has been observed as shown in the left
- Arrows show the direction
- How many equation?
- What are observations?
- How many unknown parameters?
- What are unknown parameters?
- Functional model?
- A-matrix?
- y-vector?

Exercise: GPS network

- The observations, coordinate differences, are results of the baseline processing (from phase double difference observations to coordinate differences between the points)
- Also the covariance matrices of the coordinate differences are saved in baseline processing
 Coordinate differences AV



- Coordinate differences ΔX , ΔY , ΔZ between points has been observed as shown in the left
- Arrows show the direction
- How many equation?
- What are observations?
- How many unknown parameters?
- What are unknown parameters?
- Functional model?
- A-matrix?
- y-vector?

Non-linear functional models, trilateration



Least squares estimate, BLUE, Maximum likelihood

- Least squares estimation
 - No assumptions of the probability distribution of vector of observations
 - Based on the minimizing the quadratic form $(Ax y)^T P(Ax y)$
- LSQ estimate is BLUE (Best Linear Unbiased Estimatation) if
 - Linear: LSQ estimate is linear $x = (A^T P A)^{-1} A^T P y$
 - Unbiased: LSQ estimate is unbiased $E(\hat{x}) = x$ for $\forall x$
 - Best: the variance of estimated \hat{x} is minimum when $P = \sigma_0^2 \Sigma^{-1}$
- ML estimate is BLUE if the probability distribution of observation is $y \sim N(Ax, \Sigma_y)$ and
- ML estimate is LSQ if it is BLUE and $P = \sigma_0^2 \Sigma^{-1}$

LSQ estimate is ortogonal projection



$$A = \begin{pmatrix} 1 & 0 \\ -1 & 1 \\ 0 & -1 \end{pmatrix} \quad y = \begin{pmatrix} \Delta H_{12} \\ \Delta H_{23} \\ \Delta H_{31} \end{pmatrix}$$

The columns of A matrix span the two dimensional space. Estimated \hat{y} is in in this space



LSQ is orthogonal projection

$$\hat{x} = (A^T A)^{-1} A^T y$$

$$\hat{y} = A\hat{x} = A(A^T A)^{-1} A^T y$$

$$\hat{y} = A\hat{x} = A(A^T A)^{-1} A^T y$$

$$\hat{e} = y - \hat{y} = y - A\hat{x} = y - A(A^T A)^{-1} A^T y = (I - A(A^T A)^{-1} A^T) y$$

$$\hat{y}^T \hat{e} = 0$$

$$\hat{y} \perp \hat{e}$$

Least squares process



Examples of functional models

$$\begin{split} \Phi_{i}^{k}(t) &= \rho_{i}^{k}(t) \times \frac{f}{c} + \left(h^{k}(t) - h_{i}(t)\right) \times f + ion_{i}^{k}(t) + trop_{i}^{k}(t) - N_{i}^{k} + \varepsilon & \text{GPS phase observation} \\ \tau_{obs} &= -\frac{1}{c} b_{i} \cdot Y \cdot X \cdot U \cdot N \cdot P \cdot k_{c} \\ &+ \tau_{y,abb} + \tau_{d,abb} + \tau_{rel.} \\ &+ \tau_{tides} + \tau_{o,load} + \tau_{a,load} + \tau_{h,load} \\ &+ \tau_{ion} + \tau_{instr.} + \tau_{atm,dry} + \tau_{atm,wet} + \tau_{clock} & \text{VLBI time delay} \\ \hline X_{0} + R_{a,a}(E - X_{0}) + R_{a,a}R_{\beta,e}p - X = 0 & \text{Local tie, reference point of} \\ &\alpha &= tan^{-1} \left(\frac{-sin\lambda \cdot \Delta u + cos\lambda \cdot \Delta v}{\sqrt{\Delta u^{2} + \Delta v^{2} + \Delta w^{2}}} \right) & \text{Azimuth, elevation angle,} \\ \hline \end{array}$$

$$s = \sqrt{\Delta u^2 + \Delta v^2 + \Delta w^2}$$

Litterature

- Kallio 1998:Tasoituslasku
- Cooper 1987: Control Surveys in Civil Engineering
- Leick 1995:GPS Satellite Surveying
- Hirvonen 1965: Tasoituslasku
- Mikhail 1976: Observations and Least Squares
- Teunissen 2003: Adjustment theory an introduction