# Advanced probabilistic methods Lecture 5: Mixture models and EM

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- Gaussian mixture models (GMMs)
- EM algorithm
- EM for Gaussian mixture models
- Suggested reading: Bishop: *Pattern Recognition and Machine Learning* 
  - p. 110-113 (2.3.9): Mixtures of Gaussians
  - simple\_example.pdf
  - p. 430-443: EM for Gaussian mixtures

- Standard Gaussian model (left) gives bad fit to data with clusters
- Combination of two Gaussians (right) is much better



#### Gaussian mixture models

• Gaussian mixture model with K components has density

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k N(\mathbf{x} | \mu_k, \Sigma_k).$$

- $N(x|\mu_k, \Sigma_k)$  is a **component** with its own mean  $\mu_k$  and covariance  $\Sigma_k$ .
- $\pi_k$  are the **mixing coefficients**, which satisfy  $\sum_k \pi_k = 1$ ,  $0 \le \pi_k \le 1$ .



### GMMs, latent variable representation (1/2)

- Equivalent formulation is obtained by defining **latent variables**  $\mathbf{z}_n = (z_{n1}, \dots, z_{nK})$  which tell the component for observation  $\mathbf{x}_n$
- In detail z<sub>n</sub> is a vector with exactly one element equal to 1 and other elements equal to 0. z<sub>nk</sub> = 1 means that the observation x<sub>n</sub> belongs to component k.

$$\mathbf{z}_n = (0, \dots, 0, \underbrace{1}_{k^{th} \text{ elem.}}, 0, \dots, 0)^T$$



# GMMs, latent variable representation (2/2)

• Define

$$p(z_{nk}=1)=\pi_k$$
 and  $p(\mathbf{x}_n|z_{nk}=1)=N(\mathbf{x}_n|\mu_k,\Sigma_k),$  or equivalently

$$p(\mathbf{z}_n) = \prod_{k=1}^{K} \pi_k^{z_{nk}}$$
 and  $p(\mathbf{x}_n | \mathbf{z}_n) = \prod_{k=1}^{K} N(\mathbf{x}_n | \mu_k, \Sigma_k)^{z_{nk}}$ 

Then

$$p(\mathbf{x}_n) = \sum_{\mathbf{z}_n} p(\mathbf{z}_n) p(\mathbf{x}_n | \mathbf{z}_n) = \sum_k \pi_k N(\mathbf{x}_n | \mu_k, \Sigma_k)$$

 $\rightarrow \mathbf{x}_n$  has marginally the Gaussian mixture model distribution.



 Posterior probability p(z<sub>nk</sub> = 1|x<sub>n</sub>) that observation x<sub>n</sub> was generated by component k

$$\begin{split} \gamma(z_{nk}) &\equiv p(z_{nk} = 1 | \mathbf{x}_n) = \frac{p(z_{nk} = 1)p(\mathbf{x}_n | z_{nk} = 1)}{\sum_{j=1}^{K} p(z_{nj} = 1)p(\mathbf{x}_n | z_{nj} = 1)} \\ &= \frac{\pi_k N(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^{K} \pi_j N(\mathbf{x}_n | \mu_j, \Sigma_j)} \end{split}$$

 γ(z<sub>nk</sub>) can be viewed as the responsibility that component k takes
 for explaining the observation x<sub>n</sub>

# GMM: responsibilities (2/2)

- (left) samples from a joint distribution p(z)p(x|z), showing both cluster labels z and observations x (complete data)
- (center) samples from the marginal distribution  $p(\mathbf{x})$  (incomplete data)
- (right) responsibilities of the data points, computed using *known* parameters  $\pi = (\pi_1, ..., \pi_K)$ ,  $\mu = \mu_1, ..., \mu_K$ ,  $\Sigma = (\Sigma_1, ..., \Sigma_K)$ .
- Problem: in practice  $\pi$ ,  $\mu$ , and  $\Sigma$  are usually *unknown*.



 Let X denote the observed data, and θ model parameters. The goal in maximum likelihood is to find θ:

$$\widehat{\theta} = \arg\max_{\theta} \left\{ \log p(X|\theta) \right\}$$

• If model contains latent variables Z, the log-likelihood is given by

$$\log p(X| heta) = \log \left\{ \sum_Z p(X, Z| heta) 
ight\}$$
 ,

which may be difficult to maximize analytically

• Possible solutions: 1) numerical optimization, 2) the EM algorithm (expectation-maximization)

# Idea of the EM algorithm (2/2)

- X: observed data, Z: unobserved latent variables
- {X, Z}: complete data, X: incomplete data
- In EM algorithm, we assume that the complete data log-likelihood:

$$\log p(X, Z|\theta)$$

is easy to maximize.

- Problem: Z is not observed
- Solution: maximize

$$Q(\theta, \theta_0) \equiv E_{Z|X, \theta_0} \left[ \log p(X, Z|\theta) \right]$$
$$= \sum_{Z} p(Z|X, \theta_0) \log p(X, Z|\theta)$$

where  $p(Z|X, \theta_0)$  is the posterior distribution of the latent variables computed using the current parameter estimate  $\theta_0$ 

### Illustration of the EM algorithm for GMMs



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### EM algorithm in detail

**Goal**: maximize log  $p(X|\theta)$  w.r.t.  $\theta$ 

• Initialize  $\theta_0$ 

**2 E-step** Evaluate  $p(Z|X, \theta_0)$ , and then compute

$$Q(\theta, \theta_0) = E_{Z|X, \theta_0} \left[ \log p(X, Z|\theta) \right] = \sum_{Z} p(Z|X, \theta_0) \log p(X, Z|\theta)$$

M-step Evaluate θ<sup>new</sup> using

$$\theta^{new} = \arg \max_{\theta} Q(\theta, \theta_0).$$

Set  $\theta_0 \leftarrow \theta^{new}$ 

Repeat E and M steps until convergence



Figure: 11.16 in Murphy (2012)

- $Q(\theta, \theta_0)$  is a lower bound of the log-likelihood log  $p(x|\theta)$  (Bishop, Ch. 9.4)
- EM iterates between 1) updating the lower bound (E-step), 2) maximizing the lower bound (M-step).

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- In general, Z does not have to be discrete, just replace the summation in Q(θ, θ<sub>0</sub>) by integration.
- EM-algorithm can be used to compute the MAP (maximum a posteriori) estimate by maximizing in the M-step Q(θ, θ<sub>0</sub>) + log p(θ).
- In general, EM-algorithm is applicable when the observed data X can be augmented into complete data {X, Z} such that log p(X, Z|θ) is easy to maximize; Z does not have to be latent variables but can represent, for example, unobserved values of missing or censored observations.

• Consider N independent observations  $\mathbf{x} = (x_1, \dots, x_N)$  from a two-component mixture of univariate Gaussians

$$p(x_n|\theta) = \frac{1}{2}N(x_n|0,1) + \frac{1}{2}N(x_n|\theta,1).$$
(1)

- One unknown parameter,  $\theta$ , the mean of the second component.
- Goal: estimate

$$\widehat{ heta} = rg\max_{ heta} \left\{ \log p(\mathbf{x}| heta) 
ight\}.$$

simple\_example.pdf and simple\_em.m

# EM algorithm for GMMs

$$p(\mathbf{x}) = \sum_{k=1}^{K} \pi_k N(\mathbf{x}|\mu_k, \Sigma_k)$$

- Initialize parameter μ<sub>k</sub>, Σ<sub>k</sub> and mizing coefficients π<sub>k</sub>. Repeat until convergence:
- **2** E-step: Evaluate the responsibilities using current parameter values

$$\gamma(\mathbf{z}_{nk}) = \frac{\pi_k N(\mathbf{x}_n | \mu_k, \Sigma_k)}{\sum_{j=1}^K \pi_k N(\mathbf{x}_n | \mu_k, \Sigma_j)}$$

**In M-step:** Re-estimate the parameters using the current responsibilities

$$\mu_k^{new} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) \mathbf{x}_n$$
  
$$\Sigma_k^{new} = \frac{1}{N_k} \sum_{n=1}^N \gamma(z_{nk}) (\mathbf{x}_n - \mu_k^{new}) (\mathbf{x}_n - \mu_k^{new})^T$$
  
$$\pi_k^{new} = \frac{N_k}{N}$$

- In the **M-step** the formulas for  $\mu_k^{new}$  and  $\Sigma_k^{new}$  are obtained by differentiating the expected complete data log-likelihood  $Q(\theta, \theta_0)$  with respect to the particular parameters, and setting the derivatives to zero.
- The formula for  $\pi_k^{new}$  can be derived by maximizing  $Q(\theta, \theta_0)$  under the constraint  $\sum_{k=1}^{K} \pi_k = 1$ . This can be done using the Lagrange multipliers.

### EM for GMM, caveats

- EM converges to a local optimum. In fact, the ML estimation for GMMs is not well-defined due to **singularities**: if  $\sigma_k \rightarrow 0$  for components k with a single data point, likelihood goes to infinity (fig). Remedy: prior on  $\sigma_k$ .
- **Label-switching**: non-identifiability due to the fact that cluster labels can be switched and likelihood remains the same.
- In practice it is recommended to initialize the EM for the GMM by k-means.



# GMM vs. k-means

• "Why use GMMs and not just k-means?"



- Olusters can be of different sizes and shapes
- Probabilistic assignment of data items to clusters
- Possibility to include prior knowledge (structure of the model/prior distributions on the parameters)

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- Definition of the Gaussian mixture model
- Representing the Gaussian mixture model using discrete latent variables, which specify the components (or clusters) of the observations
- ML-estimation of GMMs can be done using numerical optimization or the EM algorithm.
- The main idea of the EM algorithm is to maximize the expectation of the complete data log-likelihood, where the expectation is computed over the current posterior distribution of the latent variables.